

# Extracting Randomness Using Few Independent Sources

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## Abstract

In this work we give the first deterministic extractors from a constant number of weak sources whose entropy rate is less than  $1/2$ . Specifically, for every  $\delta > 0$  we give an explicit construction for extracting randomness from a constant (depending polynomially on  $1/\delta$ ) number of distributions over  $\{0, 1\}^n$ , each having min-entropy  $\delta n$ . These extractors output  $n$  bits, which are  $2^{-n}$  close to uniform. This construction uses several results from additive number theory, and in particular a recent one by Bourgain, Katz and Tao [BKT03] and of Konyagin [Kon03].

We also consider the related problem of constructing randomness dispersers. For any constant output length  $m$ , our dispersers use a constant number of *identical* distributions, each with min-entropy  $\Omega(\log n)$  and outputs *every* possible  $m$ -bit string with positive probability. The main tool we use is a variant of the “stepping-up lemma” used in establishing lower bound on the Ramsey number for hypergraphs (Erdős and Hajnal, [GRS80]).

**Keywords:** Randomness extractors, Ramsey graphs, Sum-Product Theorem

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# 1 Introduction

## 1.1 Background

Randomness is prevalent in computer science, and is widely used in algorithms, distributed computing, and cryptography. Perhaps the main motivation and justification for the use of randomness in computation is that randomness does exist in nature, and thus it is possible to sample natural phenomena (such as tossing coins) in order to make random choices in computation. However, there is a discrepancy between the type of random input that we expect when designing randomized algorithms and protocols, and the type of random data that can be found in nature. While randomized algorithms and protocols expect a stream of independent uniformly distributed random bits, this is too much to hope for from samples of natural phenomena.

Thus, a natural and widely studied problem has been the problem of constructing *randomness extractors*.<sup>1</sup> Loosely speaking, a *randomness extractor* is a (deterministic polynomial-time computable) function  $\text{Ext} : \{0, 1\}^n \rightarrow \{0, 1\}^m$ , such that whenever  $\mathcal{X}$  is a “good” random variable (where the definition for a “good” will be discussed shortly), then  $\text{Ext}(\mathcal{X})$  is statistically close to the uniform distribution. The random variable  $\mathcal{X}$  is supposed to model the natural data, while the output of the extractor will be used to make random choices in a probabilistic algorithm or protocol.

Intuitively, the “good” distributions should be the distributions that contain more than  $m$  bits of entropy. Otherwise, information theoretic considerations show that such an extractor cannot exist, even if it is supposed to extract randomness from a fixed source that is known to the designer. Actually, past works have converged to the measure of *min-entropy* [CG85, Zuc91] which is a stronger notion than standard (Shannon) entropy.<sup>2</sup> Indeed, the same information theoretic considerations show that if  $\text{Ext}(\mathcal{X})$  is statistically close to being uniform, the distribution  $\mathcal{X}$  must be statistically close to having min-entropy at least  $m$ . Thus possessing high min-entropy is a minimal requirement from the input distribution.

## 1.2 Seeded and Seedless Extractors

Unfortunately, it is not hard to see that there is no *single* function  $\text{Ext}$  which will produce a (close to) uniform output on *every* input distribution having high min-entropy. Previous works have dealt with this problem in two ways: the first way is to add a short truly random *seed* as a secondary input to the extractor, and the second way is to use no seed, but make further assumptions on the structure of the weak sources (in addition to the minimal assumption of it containing sufficient min-entropy).

**Seeded extractors.** One approach has been to allow the extractor to be *probabilistic*. That is, in addition to its input  $\mathcal{X}$ , we allow the function  $\text{Ext}$  to have an additional input  $\mathcal{Y}$  which is uniformly distributed. To avoid trivialities, such as the extractor that simply outputs  $\mathcal{Y}$ , we will require the input  $\mathcal{Y}$  to be (much) *shorter* than the output length  $m$ . We call the additional input  $\mathcal{Y}$

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<sup>1</sup>We note that while this has been the original motivation for studying randomness extractors, such constructs have found numerous other applications in Computer Science, and are fundamental and interesting objects in their own right.

<sup>2</sup>The *min-entropy* of a random variable  $\mathcal{X}$ , denoted by  $H^\infty(\mathcal{X})$ , is equal to  $\min_{x \in \text{Supp}(\mathcal{X})} (-\log \Pr[\mathcal{X} = x])$ . It can be easily seen that the min-entropy of a random variable  $\mathcal{X}$  is always smaller than the (Shannon) entropy of  $\mathcal{X}$ . It is known that any random variable  $\mathcal{X}$  has min-entropy  $k$  a convex combination of random variables  $\mathcal{X}_1, \dots, \mathcal{X}_t$ , where each  $\mathcal{X}_i$  is the uniform distribution over some subset of size  $2^k$ . Such random variables are called *flat*.

the *seed* of the extractor, and thus we call such constructions *seeded* extractors.<sup>3</sup> Seeded extractors have been studied extensively in the past two decades (see the survey [Sha02] and there references therein) with a series of exciting results, techniques, and applications. The current state of the art is a construction of extractors that are nearly optimal in the sense that they use a seed  $\mathcal{Y}$  of length  $O(\log n)$  bits, extracting essentially all the min-entropy of the source [LRVW03]. This in particular means that using such extractors, together with enumeration over all possible seed values, it is possible to simulate any probabilistic algorithm with polynomial overhead, using only a high min-entropy source.

**Seedless extractors.** There are many reasons to try and have seedless extractors. One is that the polynomial overhead implied by the seed enumeration above is too expensive. Another is that certain applications in computer science, such as cryptography and distributed computing, intrinsically prohibit such enumeration. For example, when using a weak source of randomness to choose an encryption key via a seeded extractor, it will certainly be *insecure* to enumerate all secret keys produced using all seeds and then send the encryptions of a secret message using all these keys. More generally, it seems that we cannot always use directly seeded extractors in cryptography (see [MP90, DS02, BST03] for different cryptographic models under weak sources of randomness).

There have been many works constructing such seedless extractors (that work for specific families of high min-entropy sources). The first to consider this problem (and indeed, the first to consider the problem of randomness extraction) was von Neumann, who gave a seedless extractor from a stream of *biased* but independent bits [vN51] (see also [Per92]). Other works, such as [Blu84, SV84, CG85, CW89, CGH<sup>+</sup>85, MU02, KZ03, TV00] constructed seedless extractors for more general families of sources.<sup>4</sup>

### 1.3 Seedless extractors from few independent sources

When seedless extraction from one source is impossible, it is natural to consider doing so from several independent sources of the same quality. After all, assuming we have one such source in nature does not seem much weaker than assuming we have several.

The first to consider this problem were Santha and Vazirani [SV84], who showed how to use  $O(\log n)$  independent “semi-random”<sup>5</sup> sources of length  $n$  and min-entropy  $\delta n$  for every constant  $\delta > 0$ .

Chor and Goldreich [CG85] were the first to consider general min-entropy sources, and proved that if  $\delta > n/2$  then two sources suffice: indeed the Hadamard-Sylvester matrix  $H : \{0, 1\}^n \times \{0, 1\}^n \rightarrow \{0, 1\}$  defined by  $H(x, y) = \langle x, y \rangle$  (with the inner product in  $\text{GF}(2)$ ) is such an extractor<sup>6</sup>. Vazirani [Vaz85] extended this to show that one can use a similar function to output a linear number of bits which are exponentially close to the uniform distribution.

Improving [CG85] seems hard. Even its disperser version, namely having a non-constant output for any two sources of entropy rate below  $1/2$ , is the notorious “Bipartite Ramsey” problem, which

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<sup>3</sup>We remark that in most of the literature, the name randomness extractor (without any qualifiers) refers to what we call here a *seeded* randomness extractor.

<sup>4</sup>The work of Trevisan and Vadhan [TV00] differs from all the rest, as well as from ours, in that it works in the computational setting; the restriction on the family of sources is computational - they are efficiently sampleable, and the extractors work assuming an unproven computational assumption.

<sup>5</sup>We will not define them formally here, only note that they are weaker than high min-entropy sources, but nevertheless one cannot extract seedlessly from only one such source.

<sup>6</sup>For this construction the  $n/2$  entropy bound is tight - there are two sources of entropy exactly  $n/2$  on which  $H$  is constant

is still open. A slight improvement for this problem was made this year by Pudlak and Rödl [PR04], who lowered the min-entropy requirement for such a disperser to  $n/2 - \sqrt{n}$ , but getting a constant  $\delta < 1/2$  remains a barrier for 2 sources. (Very recently, Barak, Kindler, Shaltiel, Sudakov and Wigderson [BKS<sup>+</sup>05] used the results of the current work to overcome this barrier and obtain such a disperser for entropy  $\delta n$  with  $\delta > 0$  an arbitrarily small constant, see Section 5.)

An alternative 2-source function is the Paley matrix,  $P : GF(p) \times GF(p) \rightarrow \{\pm 1\}$  defined by  $P(x, y) = \chi_2(x + y)$  (where  $p$  is an  $n$ -bit prime and  $\chi_2 : GF(p) \rightarrow \{\pm 1\}$  is the quadratic character).  $P$  too is an extractor with exponentially small error for entropy  $> n/2$ , and is conjectured to have the same property for entropy  $\delta n$  for all  $\delta > 0$ . While generally believed, proving this conjecture seems beyond current techniques.<sup>7</sup> Assuming even more, Zuckerman [Zuc90] showed that if this conjecture holds for *all* multiplicative characters (not just  $\chi_2$ ), then a constant (actually *poly*( $1/\delta$ )) sources suffice for extraction of linearly many bits with exponential error. The extractor we use here is *exactly* the same as Zuckerman’s (but our analysis uses no unproven assumptions).

## 1.4 Our Results

### 1.4.1 Multiple-Sample Extractors.

In this work we will be interested in extracting randomness from a *constant* (larger than 2) number of samples from a high min-entropy source.<sup>8</sup> Loosely speaking, our main result is an efficient construction to extract (almost) uniformly distributed bits from a constant number of samples from independent distributions over  $\{0, 1\}^n$  each having min-entropy at least  $\delta n$ , for an arbitrarily small constant  $\delta > 0$ . The statistical distance of our output from the uniform distribution is exponentially low (i.e.,  $2^{-\Omega(n)}$ ).<sup>9</sup> More formally, our main theorem is the following:

**Theorem 1.1** (Multiple-sample extractor). *For every constant  $\delta > 0$  there exists a constant  $\ell = (1/\delta)^{O(1)}$  and a polynomial-time computable function  $\text{Ext} : \{0, 1\}^{n \cdot \ell} \rightarrow \{0, 1\}^n$  such that for every independent random variables  $\mathcal{X}_1, \dots, \mathcal{X}_\ell$  over  $\{0, 1\}^n$  satisfying  $H^\infty(\mathcal{X}_i) \geq \delta n$  for  $i = 1, \dots, \ell$ , it holds that*

$$\text{dist}\left(\text{Ext}(\mathcal{X}_1, \dots, \mathcal{X}_\ell), U_n\right) < 2^{-\Omega(n)}$$

$U_n$  denotes the uniform distribution over the set  $\{0, 1\}^n$  and  $\text{dist}(\mathcal{X}, \mathcal{Y})$  denotes the *statistical distance* of two distributions  $\mathcal{X}$  and  $\mathcal{Y}$ . That is,

$$\text{dist}(\mathcal{X}, \mathcal{Y}) \stackrel{\text{def}}{=} \frac{1}{2} \sum_{i \in \text{Supp}(\mathcal{X}) \cup \text{Supp}(\mathcal{Y})} \left| \Pr[\mathcal{X} = i] - \Pr[\mathcal{Y} = i] \right| .$$

As mentioned in Footnote 2, we denote by  $H^\infty(\mathcal{X})$  the min-entropy of the random variable  $\mathcal{X}$ .

**Remark 1.2.** As is noted in Section 3.5, we can actually reduce the statistical distance of our extractor by using more independent samples. In particular, for every  $c$ , we can ensure that the output of our extractor will be  $2^{-cn}$ -close to the uniform distribution by multiplying the number of samples we use by a factor of  $O(c)$ . Note that if a distribution  $\mathcal{Y}$  over  $\{0, 1\}^n$  is  $2^{-cn}$ -close to the uniform distribution then in particular for every  $y \in \{0, 1\}^n$ ,  $\Pr[\mathcal{Y} = y] \in (2^{-n} - 2^{-cn}, 2^{-n} + 2^{-cn})$  (i.e.,  $\mathcal{Y}$  is even close to the uniform distribution in  $L_\infty$  norm).

<sup>7</sup>We note that it is known (using Weil’s estimate of character sums) that  $P$  is an extractor for two sources where one of them has entropy  $> n/2$  and the other one only entropy  $> \log n$ .

<sup>8</sup>We note that our results can be easily generalized to the case of a larger (i.e., super-constant) number of samples.

<sup>9</sup>Note that we get statistical distance that is much better than what is possible to obtain with seeded extractors, which cannot do better than a polynomially small distance when using a logarithmic-length seed. This low statistical distance is important for cryptographic applications.

### 1.4.2 Dispersers.

We also present a construction of *dispersers* that works for much lower entropy samples of the *same* source. A *disperser* is a relaxed variant of an extractor, which does not need to output a distribution statistically close to uniform, but rather only a distribution with large support. Thus, when talking about dispersers, it is natural to consider only the support of their input, and so talk about the inputs as sets rather than random variables (where a set of size  $2^k$  corresponds to a random variable of min-entropy  $k$ ). The formal definition of dispersers is in the next section. Our main result regarding dispersers is the following theorem:

**Theorem 1.3** (Multiple-sample same-source disperser). *There exists constants  $\ell$  and  $d$  such that for every  $m$ , there is polynomial-time computable function  $\text{Disp} : \{0, 1\}^{n^\ell} \rightarrow \{0, 1\}^m$  satisfying that for every subset  $\mathcal{X} \subseteq \{0, 1\}^n$  with  $|\mathcal{X}| \geq n^{d2^m}$*

$$\text{Disp}(\mathcal{X}, \dots, \mathcal{X}) = \{0, 1\}^m$$

*That is, if  $|\mathcal{X}| \geq n^{d2^m}$  then for every  $y \in \{0, 1\}^m$ , there exist  $x_1, \dots, x_\ell \in \mathcal{X}$  such that  $\text{Disp}(x_1, \dots, x_\ell) = y$ .*

This is a better disperser than the multiple-source extractor mentioned above in the sense that it works for very low min-entropy (indeed note that if we let  $m$  be a constant, then sets of size  $n^{O(1)}$  correspond to distributions of min-entropy of  $O(\log n)$ ). However, it still has two drawbacks: one drawback is that its output is much smaller than the input entropy (although it can still be much larger than the number of samples); another drawback is that it requires all its input samples to come from the *same* distribution  $\mathcal{X}$  (rather than from different distributions, as in our extractor). We consider the second drawback to be the more serious one. The construction of this disperser closely follows the “stepping-up” technique of Erdős and Hajnal [GRS80, Sec. 4.7] for giving an (explicit) lower-bound on the Ramsey number of hypergraphs. We remark that we work for worse entropy than this lower-bound<sup>10</sup> because we want to output a super-constant number of bits (that does not depend on the number of samples).

## 1.5 Our Techniques

**An Erdős-Szemerédi theorem for finite fields.** Our main tools for Theorem 1.1 are several results from additive number theory and in particular a relatively recent result by Bourgain, Katz and Tao [BKT03]. They proved an analog of the Erdős-Szemerédi [ES83] theorem for finite prime fields. Let  $\mathcal{A}$  be a subset of some field  $\mathbb{F}$ . We define the set  $\mathcal{A} + \mathcal{A}$  to equal  $\{a + b \mid a, b \in \mathcal{A}\}$  and the set  $\mathcal{A} \cdot \mathcal{A}$  to equal  $\{a \cdot b \mid a, b \in \mathcal{A}\}$ . Note that  $|\mathcal{A}| \leq |\mathcal{A} + \mathcal{A}| \leq |\mathcal{A}|^2$  (and similarly  $|\mathcal{A}| \leq |\mathcal{A} \cdot \mathcal{A}| \leq |\mathcal{A}|^2$ ). An example for a set  $\mathcal{A}$  where  $\mathcal{A} + \mathcal{A}$  is small (of size about  $2|\mathcal{A}|$ ) is an *arithmetic progression*. An example for a set  $\mathcal{A}$  where  $\mathcal{A} \cdot \mathcal{A}$  is small is a *geometric progression*. The Erdős-Szemerédi theorem is that for every set  $\mathcal{A} \subseteq \mathbb{N}$ , either  $\mathcal{A} + \mathcal{A}$  or  $\mathcal{A} \cdot \mathcal{A}$  is of size at least  $|\mathcal{A}|^{1+\epsilon_0}$ , for some universal constant  $\epsilon_0$ . In some sense, one can view this theorem as saying that a set of integers can’t be simultaneously close to both an arithmetic progression and a geometric progression.

A natural question, with many diverse applications in geometry and analysis, is whether this theorem also holds in *finite* fields. A first observation is that this theorem is *false* in a field  $\mathbb{F}$  that contains a non-trivial subfield  $\mathbb{F}'$ . This because if we let  $\mathcal{A} = \mathbb{F}'$  then  $\mathcal{A} + \mathcal{A} = \mathcal{A} \cdot \mathcal{A} = \mathcal{A}$ .

<sup>10</sup>Using  $c$  samples these lower-bound yield a 1-bit-output disperser for  $\log^{(\Theta(c))} n$  entropy, where we use  $\log^{(i)} n$  to denote iterating  $\log$   $i$  times on  $n$ .

However, [BKT03] showed that a variant of this theorem does hold in a finite field with no non-trivial subfields.<sup>11</sup> In particular it holds in the fields  $\text{GF}(p)$  and  $\text{GF}(2^p)$  for every prime  $p$ . That is, they proved the following:

**Theorem 1.4** ([BKT03]). *Let  $\delta > 0$  be some constant and let  $\mathbb{F}$  be a field with no subfield  $\mathbb{F}'$  of size in  $[|\mathbb{F}|^{\delta/2}, |\mathbb{F}|)$ . Let  $\mathcal{A} \subseteq \mathbb{F}$  be a set such that  $|\mathcal{A}| \in (|\mathbb{F}|^\delta, |\mathbb{F}|^{1-\delta})$ . Then, there exist some constant  $\epsilon$  (depending on  $\delta$ ) such that*

$$\max\{|\mathcal{A} + \mathcal{A}|, |\mathcal{A} \cdot \mathcal{A}|\} > |\mathcal{A}|^{1+\epsilon}$$

In [BKT03], the dependence of the constant  $\epsilon$  on  $\delta$  was not specified (and examining the proof sees that it is probably of the form  $\epsilon = 2^{-\Omega(1/\delta)}$ ). In Appendix A we give a proof of Theorem 1.4 (using some of their lemmas) with  $\epsilon = \Theta(\delta)$ .<sup>12</sup> Konyagin [Kon03] gave a stronger result for prime fields, and showed that, as long as  $|\mathcal{A}| < |\mathbb{F}|^{0.99}$ ,  $\epsilon$  can be made independent of the size of  $\mathcal{A}$  (even if  $|\mathcal{A}|$  is very small). That is, he proved the following theorem:

**Theorem 1.5** ([Kon03]). *There exist some constant  $\epsilon_0$  such that for every  $\mathbb{F}$  which is a field of prime order, and every  $\mathcal{A} \subseteq \mathbb{F}$  with  $|\mathcal{A}| < |\mathbb{F}|^{0.99}$ ,*

$$\max\{|\mathcal{A} + \mathcal{A}|, |\mathcal{A} \cdot \mathcal{A}|\} > |\mathcal{A}|^{1+\epsilon_0}$$

## 1.6 How is this related to extractors?

**Using Theorem 1.5 to obtain dispersers.** Theorem 1.5 actually already implies some kind of multiple-sample disperser that was not previously known. This is because it implies the following corollary:

**Corollary 1.6.** *There exist some constant  $\epsilon_0$  such that for every  $\mathbb{F}$  which is a field of prime order, and every  $\mathcal{A} \subseteq \mathbb{F}$  with  $|\mathcal{A}| < |\mathbb{F}|^{0.99}$ ,*

$$|\mathcal{A} \cdot \mathcal{A} + \mathcal{A}| > |\mathcal{A}|^{1+\epsilon_0}$$

In other words, it implies that  $|f(\mathcal{A}, \mathcal{A}, \mathcal{A})| \geq |\mathcal{A}|^{1+\epsilon_0}$  where  $f(a, b, c) = a \cdot b + c$ . Yet this implies that for every set  $\mathcal{A}$  with  $|\mathcal{A}| \geq |\mathbb{F}|^\delta$ , if we compose  $f$  with itself  $O(\log(1/\delta))$  times we get a function  $g : \mathbb{F}^\ell \rightarrow \mathbb{F}$  (for  $\ell = 2^{O(\log(1/\delta))} = (1/\delta)^{O(1)}$ ) such that  $g(\mathcal{A}, \dots, \mathcal{A}) = \mathbb{F}$  (where  $g(\mathcal{A}, \dots, \mathcal{A})$  denotes the image of  $g$  on  $\mathcal{A}^\ell$ ).<sup>13</sup> If we identify the field  $\mathbb{F}$  with the strings  $\{0, 1\}^n$ , then we see that this function  $g$  (which is obviously polynomial-time computable) is already some kind of a disperser.

**Obtaining extractors.** To obtain an *extractor*, rather than a disperser we would like to obtain a *statistical* analog of Theorem 1.5. Consider the following notation: if  $\mathcal{X}$  is a random variable then we define  $\mathcal{X} + \mathcal{X}$  to be the distribution obtained by choosing  $a, b$  independently at random from  $\mathcal{X}$  and outputting  $a + b$ , and define  $\mathcal{X} \cdot \mathcal{X}$  in a similar way. Then, a statistical analog of Theorem 1.5 would be that there exists some  $\epsilon > 0$  such that for every random variable  $\mathcal{X}$  with min-entropy at most  $0.9 \log |\mathbb{F}|$ , either the distribution  $\mathcal{X} + \mathcal{X}$  or the distribution  $\mathcal{X} \cdot \mathcal{X}$  has min-entropy at least

<sup>11</sup>By a *trivial* subfield in this context we mean a subfield that is either the entire field or very small (i.e., of constant size).

<sup>12</sup>The proof of Appendix A also has the advantage of proving the theorem for all fields  $\mathbb{F}$  simultaneously. [BKT03] prove Theorem 1.4 for prime fields, and then only mention how it can be generalized to other fields.

<sup>13</sup>This is because each time we apply  $f$  we grow in the set size from  $m$  to  $m^{1+\epsilon_0}$ . We note that one needs to analyze separately the case that  $|\mathcal{A}| > |\mathbb{F}|^{0.99}$  but this can be done. This function  $g$  is described in more detail in the beginning of Section 3.

$(1 + \epsilon)H^\infty(\mathcal{X})$ . Unfortunately, this is false: for every prime field  $\mathbb{F}$ , there is a random variable  $\mathcal{X}$  which is uniform over some set of size  $2^k$  (for  $k \ll 0.9 \log |\mathbb{F}|$ ) and such that for both  $\mathcal{X} + \mathcal{X}$  and  $\mathcal{X} \cdot \mathcal{X}$ , a constant fraction of the probability mass is concentrated in a set of size  $O(2^k)$ .<sup>14</sup>

Even though the statistical analog for [Theorem 1.5](#) does not hold, we show that a statistical analog for [Corollary 1.6](#) *does* hold. That is, we prove (in [Lemma 3.1](#)) that there exists some constant  $\epsilon_0 > 0$  such that for every random variable  $\mathcal{X}$  over a prime field  $\mathbb{F}$ , the distribution  $\mathcal{X} \cdot \mathcal{X} + \mathcal{X}$  has (up to some negligible statistical distance) min-entropy at least  $(1 + \epsilon_0)H^\infty(\mathcal{X})$ . (In fact, we need to (and do) prove a more general statement regarding distributions of the form  $\mathcal{X} \cdot \mathcal{Y} + \mathcal{Z}$ .) The proof of this lemma is the main technical step in our extractor. The proof uses [Theorem 1.5](#), along with some other additive number-theoretic results of Ruzsa [[Ruz96](#)] and Gowers [[Gow98](#)].

We use this lemma to show that the function  $g$  sketched above is actually not just a disperser but an *extractor*. That is, we show that for every random variable  $\mathcal{X}$  of min-entropy at least  $\delta \log |\mathbb{F}|$ , the random variable  $g(\mathcal{X}, \dots, \mathcal{X})$  (where  $g$  is applied to  $\ell$  independent copies of  $\mathcal{X}$ ) not only has large support but is in fact very close to the uniform distribution. (In fact, we need to (and do) prove a stronger version of this statement. Namely, that  $g(\mathcal{X}_1, \dots, \mathcal{X}_\ell)$  is close to uniform for every independent random variables  $\mathcal{X}_1, \dots, \mathcal{X}_\ell$  with  $H^\infty(\mathcal{X}_i) \geq \delta \log |\mathbb{F}|$  for  $i = 1, \dots, \ell$ .)

## 2 Preliminaries

In this section we establish our definitions for multiple-sample extractors and dispersers. Unfortunately, such definitions tend to have a large number of parameters.

**Definition 2.1** (Multiple-sample extractor). A function  $\text{Ext} : \{0, 1\}^{n-\ell} \rightarrow \{0, 1\}^m$  is called a  $\ell$ -sample  $(k, m, \epsilon)$ -extractor if for every independent random variables  $\mathcal{X}_1, \dots, \mathcal{X}_\ell$  satisfying  $H^\infty(\mathcal{X}_i) \geq k$  for  $i = 1, \dots, \ell$  it holds that

$$\text{dist}\left(\text{Ext}(\mathcal{X}_1, \dots, \mathcal{X}_\ell), U_m\right) < \epsilon$$

**Parameters and qualifiers:**

- The parameter  $\ell$  is called the *number of samples* of the extractor. In all of our constructions we will use  $\ell$  which is a constant independent of the input length.
- The parameter  $m$  is called the *output length* of the extractor. Usually, we'll use  $m = n$ .
- The parameter  $k$  is called the min-entropy requirement by the extractor.
- The parameter  $\epsilon$  is called the statistical distance of the extractor.
- One can also make a weaker definition in which the extractor is required to output a close to uniform value only if the variables  $\mathcal{X}_1, \dots, \mathcal{X}_\ell$  are identically distributed. We call such an extractor a *same-source* extractor. Thus, we will sometimes say that a standard (as per [Definition 2.1](#)) multiple-sample extractor is a *different-source* extractor.

We now define the weaker notion of a disperser:

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<sup>14</sup>An example for such a random variable  $\mathcal{X}$  is a random variable which is with probability half an arithmetic progression and with probability half a geometric progression.

**Definition 2.2** (Multiple-sample disperser). A function  $\text{Disp} : \{0, 1\}^{n \cdot \ell} \rightarrow \{0, 1\}^m$  is called a  $\ell$ -sample  $(k, m)$ -disperser if for all sets  $\mathcal{X}_1, \dots, \mathcal{X}_\ell$  satisfying  $|\mathcal{X}_i| \geq 2^k$  for  $i = 1, \dots, \ell$  it holds that

$$\text{Disp}(\mathcal{X}_1, \dots, \mathcal{X}_\ell) = \{0, 1\}^m$$

**Notes:**

- We will use the same names for the parameters and qualifiers of dispersers as we used for extractors (e.g., number of samples, output length, same-source/different-source).
- In previous literature dispersers are usually defined with a parameter  $\epsilon$  analogous to the statistical distance of extractors, requiring that  $|\text{Disp}(\mathcal{X}_1, \dots, \mathcal{X}_\ell)| \geq (1 - \epsilon)2^m$  instead of the requirement we make. Thus, one can think of our definition as setting  $\epsilon = 0$ . However, in our particular setting of independent samples, this is essentially without loss of generality as we can convert a disperser satisfying the weaker definition with any constant  $\epsilon < 1/2$  into a disperser satisfying the stronger definition with only a constant factor increase in the number of samples.<sup>15</sup>

**Basic facts and observations.** The following facts regarding multiple-sample extractors and dispersers are either known or easy to verify:

- There does not exist a 1-sample disperser (and hence extractor) even with only one bit of output and even if the source is assumed to have  $n - 1$  bits of min-entropy.
- There is a simple construction for a 2-sample *same-source* 1-bit output extractor whenever the source min-entropy is larger than  $O(\log \frac{1}{\epsilon})$  (where  $\epsilon$  is the statistical distance). This is the extractor that on input  $x, y$  outputs 1 if  $x > y$ .
- In contrast, by Ramsey-type arguments, a 2-sample same-source disperser (and hence an extractor) with more than 1 bit of output requires the input min-entropy to be at least  $\log n$ . The same holds for a different-source 1-bit output disperser.

The best known previous explicit constructions for both cases require the min-entropy to be more than  $n/2$  [CG85, Vaz87]. Indeed, the function  $H : \{0, 1\}^{2n} \rightarrow \{0, 1\}$  defined as follows  $H(x, y) = \sum x_i y_i \pmod{2}$  (i.e., the adjacency function of the Hadamard graph) is known to be a 1-bit output extractor for sources with more than  $n/2$  entropy [CG85]. This is essentially the best known previous construction in terms of the minimum entropy requirement.<sup>16</sup> There has been some improvement in obtaining variants of this extractor that have a larger output size [Vaz87, Elb03, DO03] (although these still require at least  $n/2$  entropy).

- Using enumeration over all possible seeds, one can use a *seeded* extractor to obtain a polynomial-samples same-source extractor with the same requirement over the min-entropy. It is also possible to generalize this to work for *different sources* [SRW03].

<sup>15</sup>This can be done for example by identifying the set  $\{0, 1\}^n$  with elements of a prime field  $\mathbb{F}$  (as we do in the sequel) and using the Cauchy-Davenport theorem. This theorem says that if  $\mathcal{A}$  and  $\mathcal{B}$  are subsets of a prime field  $\mathbb{F}$ , then the set  $\mathcal{A} + \mathcal{B} = \{a + b \mid a \in \mathcal{A}, b \in \mathcal{B}\}$  is of size at least  $\min\{|\mathbb{F}|, |\mathcal{A}| + |\mathcal{B}| - 1\}$ . Hence, if  $\mathcal{A}_1, \dots, \mathcal{A}_k$  are sets of size  $\epsilon|\mathbb{F}|$ ,  $k > 1/\epsilon$  and  $|\mathcal{A}_i| > \epsilon|\mathbb{F}|$  for all  $i \in [k]$  then  $\mathcal{A}_1 + \dots + \mathcal{A}_k = \mathbb{F}$ .

<sup>16</sup>This is assuming we make the same entropy requirement from both sources. As mentioned above, it is known that the adjacency function of the Paley graph is a 1-bit output extractor which works as long as one of its inputs has more than  $n/2$  entropy while the other input only needs to have more than  $\log n$  entropy.

**Explicit constructions.** In the definition of extractors and dispersers we did not require these functions to be efficiently computable. However, we will naturally be interested in obtaining extractors and dispersers that are computable in polynomial-time. We call such constructions *explicit*.

### 3 Constructing a Multiple-Sample Extractor

In this section we prove [Theorem 1.1](#). That is, we construct an extractor  $\text{Ext}$  (where  $\text{Ext} : \{0, 1\}^{\ell \cdot n} \rightarrow \{0, 1\}^n$ ) such that  $\text{Ext}(\mathcal{X}_1, \dots, \mathcal{X}_\ell)$  is statistically close to the uniform distribution for every independent random variables  $\mathcal{X}_1, \dots, \mathcal{X}_\ell$  over  $\{0, 1\}^n$  with high enough min-entropy (i.e., at least  $\delta n$  where  $\ell = \text{poly}(1/\delta)$ ). Our extractor will be very simple, involving only a recursive composition of the operation  $(a, b, c) \mapsto a \cdot b + c$ . As noted in the introduction, this is the same exact construction used by Zuckerman [[Zuc90](#)].

Formally, we fix a suitable field  $\mathbb{F}$  (see below) and define the functions  $\text{Ext}^i$  for all  $i \in \mathbb{N}$  recursively as follows:

- For every  $i$ , the function  $\text{Ext}^i$  will be a mapping from  $\mathbb{F}^{3^i}$  to  $\mathbb{F}$ .
- $\text{Ext}^0 : \mathbb{F}^{3^0} = \mathbb{F} \rightarrow \mathbb{F}$  is the identity function  $\text{Ext}^0(x) = x$ .
- Assume  $\text{Ext}^i$  is already defined. We define  $\text{Ext}^{i+1} : \mathbb{F}^{3^{i+1}} = \left(\mathbb{F}^{3^i}\right)^3 \rightarrow \mathbb{F}$  as follows:  
for every  $x_1, x_2, x_3 \in \mathbb{F}^{3^i}$ ,

$$\text{Ext}^{i+1}(x_1, x_2, x_3) \stackrel{\text{def}}{=} \text{Ext}^i(x_1) \cdot \text{Ext}^i(x_2) + \text{Ext}^i(x_3)$$

Our extractor will be equal to  $\text{Ext}^i$  for a suitably chosen constant  $i$ . When extracting from a string in  $\{0, 1\}^n$  we will choose as the field  $\mathbb{F}$  to be of the form  $\text{GF}(N)$  for some prime  $N \in [2^n, 2^{n+1})$  (and hence we can identify any string in  $\{0, 1\}^n$  with an element in the field  $\mathbb{F}$ ). We postpone the issue of finding such a prime to [Section 3.4](#).

[Theorem 1.1](#) will follow from the following two lemmas:

**Lemma 3.1.** *There exists some constant  $\epsilon > 0$  such that for every distributions  $\mathcal{A}, \mathcal{B}, \mathcal{C}$  over a prime field  $\mathbb{F}$  each with min-entropy at least  $m$ , the distribution  $\mathcal{A} \cdot \mathcal{B} + \mathcal{C}$  is  $2^{-\epsilon m}$ -close to having min-entropy at least  $\min\{(1 + \epsilon)m, 0.9 \log |\mathbb{F}|\}$ .*

**Lemma 3.2.** *Let  $\mathcal{A}_1, \dots, \mathcal{A}_9$  be 9 independent distributions over a prime field  $\mathbb{F}$  each with min-entropy at least  $0.9 \log |\mathbb{F}|$ . Then,  $\text{Ext}^2(\mathcal{A}_1, \dots, \mathcal{A}_9)$  is of distance at most  $|F|^{-0.01}$  from the uniform distribution over  $\mathbb{F}$ .*

Lemmas [3.1](#) and [3.2](#) imply [Theorem 1.1](#). Indeed, [Lemma 3.1](#) implies that for every constant  $\delta > 0$ , if we let  $i = \log_{(1+\epsilon)}(1/\delta)$  then the output of  $\text{Ext}^i$  is close to having min-entropy  $0.9 \log |\mathbb{F}|$ . Note that the number of samples  $\text{Ext}^i$  requires is  $3^i$  which is polynomial in  $1/\delta$ . We use [Lemma 3.2](#) to get from min-entropy  $0.9 \log |\mathbb{F}|$  to a close to uniform output. Using the union bound, the statistical distance of the extractors output from the uniform distribution on  $\mathbb{F}$  will be at most  $N^{-\Omega(1)}$ . We defer the proof of [Lemma 3.2](#) to [Section 3.6](#) and now turn to proving [Lemma 3.1](#).<sup>17</sup>

<sup>17</sup>Note that we could prove [Theorem 1.1](#) without using [Lemma 3.2](#) by applying the previously known two-sample extractors that work for entropy larger than  $\frac{\log |\mathbb{F}|}{2}$ . In addition, Salil Vadhan observed that using the fact that the function family  $\{h_{b,c}\}$  (where  $h_{b,c}(a) = a \cdot b + c$ ) is pairwise independent and the Leftover Hash Lemma of [[HILL89](#)], one can prove that under the conditions of [Lemma 3.2](#) the first  $0.8 \log |F|$  bits of  $\mathcal{A}_1 \cdot \mathcal{A}_2 + \mathcal{A}_3$  are within statistical distance  $|F|^{-0.01}$  to the uniform distribution. This is also sufficient to prove [Theorem 1.1](#).

### 3.1 Basic Facts and Notations

Before proving [Lemma 3.1](#), we introduce some notations and some tools that will be used in the course of the proof. We identify a set  $\mathcal{A}$  with the uniform distribution on elements of  $\mathcal{A}$ . If  $\mathcal{D}$  is a distribution then we denote by  $\text{cp}(\mathcal{D})$  the collision probability of  $\mathcal{D}$ ; that is,  $\text{cp}(\mathcal{D}) = \Pr_{x,y \leftarrow \mathbb{R}\mathcal{D}}[x = y]$ . Note that for a set  $\mathcal{A}$ ,  $\text{cp}(\mathcal{A}) = \frac{1}{|\mathcal{A}|}$ . If  $\mathcal{A}$  and  $\mathcal{B}$  are distributions over subsets of the field  $\mathbb{F}$ , then we denote by  $\mathcal{A} + \mathcal{B}$  the *distribution* obtained by picking  $a$  at random from  $\mathcal{A}$ ,  $b$  at random from  $\mathcal{B}$  and outputting  $a + b$ . Note that this distribution may be far from the uniform distribution with the same support. We define the distribution  $\mathcal{A} \cdot \mathcal{B}$  in a similar way. For  $k \in \mathbb{Z}$  and  $\mathcal{A}$  a set or a distribution we define  $k\mathcal{A}$  and  $\mathcal{A}^k$  in the natural way.<sup>18</sup>

We say that a distribution  $\mathcal{X}$  is a convex combination of distributions  $\mathcal{X}_1, \dots, \mathcal{X}_m$  if there exist numbers  $p_1, \dots, p_m \in [0, 1]$  such that  $\sum_i p_i = 1$  and the random variable  $\mathcal{X}$  (when looked at as a vector of probabilities) is equal to  $\sum_i p_i \mathcal{X}_i$ .

The following lemmas would be useful:

**Lemma 3.3.** *Let  $\mathcal{X}$  and  $\mathcal{Y}$  be distributions, then  $\Pr[\mathcal{X} = \mathcal{Y}] \leq \sqrt{\text{cp}(\mathcal{X})\text{cp}(\mathcal{Y})} \leq \max\{\text{cp}(\mathcal{X}), \text{cp}(\mathcal{Y})\}$ .*

*Proof.* For every  $i$  in the union of the supports of  $\mathcal{X}$  and  $\mathcal{Y}$ , let  $x_i$  denote the probability that  $\mathcal{X} = i$  and let  $y_i$  denote the probability that  $\mathcal{Y} = i$ . Then

$$\Pr[\mathcal{X} = \mathcal{Y}] = \sum_i x_i y_i \leq \sqrt{\sum_i x_i^2 \sum_j y_j^2}$$

by Cauchy-Schwartz, however this is the geometric mean of  $\text{cp}(\mathcal{X})$  and  $\text{cp}(\mathcal{Y})$  (which is less than the maximum of these quantities).  $\square$

**Corollary 3.4.** *Let  $\mathcal{A}, \mathcal{C}$  be distributions over  $\mathbb{F}$ . Then  $\text{cp}(\mathcal{A} + \mathcal{C}) \leq \sqrt{\text{cp}(\mathcal{A} - \mathcal{A})\text{cp}(\mathcal{C} - \mathcal{C})}$ .*

*Proof.*

$$\text{cp}(\mathcal{A} + \mathcal{C}) = \Pr[\mathcal{A} + \mathcal{C} = \mathcal{A}' + \mathcal{C}'] = \Pr[\mathcal{A} - \mathcal{A}' = \mathcal{C} - \mathcal{C}']$$

which is smaller than  $\sqrt{\text{cp}(\mathcal{A} - \mathcal{A})\text{cp}(\mathcal{C} - \mathcal{C})}$  by [Lemma 3.3](#).  $\square$

**Lemma 3.5.** *Suppose that  $\mathcal{X}$  is a convex combination of distributions  $\mathcal{X}_1, \dots, \mathcal{X}_m$ . Then  $\text{cp}(\mathcal{X}) \leq \max\{\text{cp}(\mathcal{X}_1), \dots, \text{cp}(\mathcal{X}_m)\}$ .*

*Proof.* Using induction on  $m$ , this reduces to the case that  $m = 2$  (since we can treat the combination of  $\mathcal{X}_1, \dots, \mathcal{X}_{m-1}$  as one distribution whose collision probability is bounded using induction). However in this case  $\mathcal{X} = \alpha\mathcal{X}_1 + (1 - \alpha)\mathcal{X}_2$  and then

$$\text{cp}(\mathcal{X}) = \alpha^2 \text{cp}(\mathcal{X}_1) + 2\alpha(1 - \alpha) \Pr[\mathcal{X}_1 = \mathcal{X}_2] + (1 - \alpha)^2 \text{cp}(\mathcal{X}_2)$$

However,  $\Pr[\mathcal{X}_1 = \mathcal{X}_2] \leq \max\{\text{cp}(\mathcal{X}_1), \text{cp}(\mathcal{X}_2)\}$  by [Lemma 3.3](#) and so we're done.  $\square$

**Lemma 3.6.** *Let  $\mathcal{X}$  be a distribution such that  $\text{cp}(\mathcal{X}) \leq \frac{1}{KL}$ . Then  $\mathcal{X}$  is of statistical distance  $\frac{1}{\sqrt{L}}$  from having min-entropy at least  $\log K$ .*

<sup>18</sup>E.g., for  $k > 0$ ,  $k\mathcal{A} = \underbrace{\mathcal{A} + \dots + \mathcal{A}}_{k \text{ times}}$  and for  $k < 0$ ,  $k\mathcal{A} = \underbrace{-\mathcal{A} - \dots - \mathcal{A}}_{|k| \text{ times}}$ .

*Proof.* We can split  $\mathcal{X}$  into a convex combination  $\mathcal{X} = \alpha\mathcal{X}' + (1 - \alpha)\mathcal{X}''$ , where  $\mathcal{X}'$  contains the “big” elements of  $\mathcal{X}$  that are obtained with probability at least  $\frac{1}{K}$  and  $\mathcal{X}''$  contains the rest of  $\mathcal{X}$ . We see that  $\text{cp}(\mathcal{X}) \geq \alpha^2 \text{cp}(\mathcal{X}')$ , and so  $\frac{1}{KL} \geq \frac{\alpha^2}{K}$  and thus  $\alpha \leq \frac{1}{\sqrt{L}}$ . However,  $H^\infty(\mathcal{X}'') \geq \log K$  and so we’re done.  $\square$

Note that if  $H^\infty(\mathcal{X}) \geq k$  then clearly  $\text{cp}(\mathcal{X}) \leq 2^{-k}$ . Together with [Lemma 3.6](#) this implies that, up to an arbitrarily close to 1 multiplicative factor and an exponentially small (in the min-entropy) statistical distance,  $H^\infty(X)$  is approximately equal to  $\log(1/\text{cp}(\mathcal{X}))$ . (The quantity  $\log(1/\text{cp}(\mathcal{X}))$  is sometimes called the *2-entropy* or the *Renyi entropy* of  $\mathcal{X}$ .)

### 3.2 Additive Number-Theoretic Results

The following two lemmas hold for any Abelian group, and so the “+” operator may be replaced in them by “ $\cdot$ ”. Note that we didn’t state these lemmas with the most optimal choice of constants.

**Lemma 3.7** ([\[Ruz96\]](#) (see also [\[Nat96, Ruz89\]](#))). *Let  $\mathcal{A}, \mathcal{B}$  be subsets of some Abelian group  $G$  with  $|\mathcal{A}| = |\mathcal{B}| = M$  and let  $\rho > 0$  be some number. If  $|\mathcal{A} + \mathcal{B}| \leq M^{1+\rho}$  then  $|\mathcal{A} + \mathcal{A}| \leq M^{1+2\rho}$ .*

In other words, if  $\mathcal{A} + \mathcal{B}$  is “small” for some  $\mathcal{B}$  then  $\mathcal{A} + \mathcal{A}$  is small. Note that [Lemma 3.7](#) implies that we can replace  $\mathcal{A} + \mathcal{A}$  and  $\mathcal{A} \cdot \mathcal{A}$  in [Theorem 1.4](#) with  $\mathcal{A} + \mathcal{B}$  and  $\mathcal{A} \cdot \mathcal{B}$ .

We’ll use the following lemma of Gowers (which is a quantitative improvement of a result by Balog and Szemerédi). We state it here using different sets  $\mathcal{A}, \mathcal{B}$  (similarly to the way it is quoted in [\[BKT03\]](#), although in [\[Gow98\]](#) it is stated with  $\mathcal{A} = \mathcal{B}$ ).

**Lemma 3.8** (Proposition 12, [\[Gow98\]](#)). *Let  $\mathcal{A}, \mathcal{B}$  be subsets of some group  $G$  with  $|\mathcal{A}| = |\mathcal{B}| = M$  and let  $\rho > 0$  be such that  $\text{cp}(\mathcal{A} + \mathcal{B}) \geq \frac{1}{M^{1+\rho}}$ . Then, there exists  $\mathcal{A}' \subseteq \mathcal{A}$  and  $\mathcal{B}' \subseteq \mathcal{B}$  such that  $|\mathcal{A}'|, |\mathcal{B}'| \geq M^{1-10\rho}$  and  $|\mathcal{A}' + \mathcal{B}'| \leq M^{1+10\rho}$ .*

We remark that it can be shown that if  $\mathcal{A} = \mathcal{B}$  in the conditions of this lemma, then  $\mathcal{A}' = \mathcal{B}'$  in its conclusion. A slightly different proof of this lemma (along with an extension to sums of more than two sets) can also be found in [\[SSV04\]](#).

### 3.3 Proof of [Lemma 3.1](#)

**Fixing  $\epsilon$ .** We fix  $\epsilon$  small enough such that [Theorem 1.5](#) ensures us that if  $\mathcal{X}$  is a set of size at least  $M^{1-10^4\epsilon}$  then  $\max\{|\mathcal{X} \cdot \mathcal{X}|, |\mathcal{X} + \mathcal{X}|\}$  is at least  $M^{1+10^4\epsilon}$  (e.g., we can take  $\epsilon = \epsilon_0/10^9$ ). **Note:** This is the only place in the proof we use the fact that  $\mathbb{F}$  is a prime field. In particular, using [Theorem 1.4](#) instead of [Theorem 1.5](#)) our proof yields also a variant of [Lemma 3.1](#) for non-prime fields (see [Lemma 3.14](#)).

**Statistical analog of [Theorem 1.5](#).** Roughly speaking, [Theorem 1.5](#) says that if  $|\mathcal{A} \cdot \mathcal{A}|$  is “small” then  $|\mathcal{A} + \mathcal{A}|$  is “large”. By combining [Theorem 1.4](#) with [Lemma 3.7](#), it is possible get a “different sets” analog of this statement and show that for every set  $\mathcal{A}$  of size  $M$ , if  $|\mathcal{A} \cdot \mathcal{B}|$  is “small” (i.e.  $\leq M^{1+\epsilon}$ ) for some set  $\mathcal{B}$  of size  $M$ , then  $|\mathcal{A} + \mathcal{C}|$  is “large” ( $\geq M^{1+\epsilon}$ ) for *every* set  $\mathcal{C}$  of size  $M$ .

At this point we would have liked to obtain a collision probability analog of this statement. That is, we would like to prove that if  $\mathcal{A}, \mathcal{B}, \mathcal{C}$  are uniform distributions over sets of size  $M$ , and  $\text{cp}(\mathcal{A} \cdot \mathcal{B}) \geq \frac{1}{M^{1+\epsilon}}$  then  $\text{cp}(\mathcal{A} + \mathcal{C}) \leq \frac{1}{M^{1+\epsilon}}$ . Unfortunately, this is false, as can be witnessed by considering the following counterexample: Let  $M = N^{0.1}$  and let  $\mathcal{A} = \mathcal{B} = \mathcal{C}$  be a distribution that

is an arithmetic progression of size  $M$  with probability  $\frac{1}{2}$  and a geometric progression of size  $M$  with probability  $\frac{1}{2}$ . For such a distribution, both  $\text{cp}(\mathcal{A} + \mathcal{A})$  and  $\text{cp}(\mathcal{A} \cdot \mathcal{A})$  are at least  $\Omega(1/M)$ .<sup>19</sup>

However, we are able to prove a slightly weaker statement. That is, we show that if  $\mathcal{A} \cdot \mathcal{B}$  is small in *set-size*, then  $\mathcal{A} + \mathcal{C}$  shrinks significantly in *collision probability*. This is stated in the following lemma:

**Lemma 3.9.** *Let  $\mathcal{A} \subseteq \mathbb{F}$  with  $|\mathcal{A}| \geq M^{1-10\epsilon}$  such that  $|\mathcal{A} \cdot \mathcal{B}| \leq M^{1+\epsilon}$ , for some  $\mathcal{B}$  with  $|\mathcal{B}| \geq M^{1-10\epsilon}$ . Then  $\text{cp}(\mathcal{A} + \mathcal{C})$  is smaller than  $\frac{1}{M^{1+10\epsilon}}$  for all sets  $\mathcal{C}$  of size  $M$ .*

*Proof.* If  $\text{cp}(\mathcal{A} + \mathcal{C}) \geq \frac{1}{M^{1+10\epsilon}}$  then by applying Lemma 3.8 (with  $\rho = 20\epsilon$ ) it holds that there exists subsets  $\mathcal{A}', \mathcal{C}'$  of  $\mathcal{A}$  and  $\mathcal{C}$  respectively such that  $|\mathcal{A}'|, |\mathcal{C}'| \geq M^{1-200\epsilon}$  but  $|\mathcal{A}' + \mathcal{C}'| \leq M^{1+200\epsilon}$ . This means by Lemma 3.7 that  $|\mathcal{A}' + \mathcal{A}'| \leq M^{1+2000\epsilon}$ . However, this means that  $|\mathcal{A}' \cdot \mathcal{A}'| \geq M^{1+10^4\epsilon}$  by Theorem 1.5 which implies that  $|\mathcal{A}' \cdot \mathcal{B}| \geq M^{1+10^3\epsilon}$  by Lemma 3.7. However, since  $\mathcal{A}' \subseteq \mathcal{A}$  this implies that  $|\mathcal{A} \cdot \mathcal{B}| \geq M^{1+10^3\epsilon}$ , contradicting our initial assumption.  $\square$

We call a set  $\mathcal{A}$  satisfying the conclusion of Lemma 3.9 *plus-friendly* (since its collision probability grows when it is added with any other set of the same size). Reversing the roles of addition and multiplication we obtain

**Lemma 3.10.** *Let  $\mathcal{A} \subseteq \mathbb{F}$  with  $|\mathcal{A}| \geq M^{1-10\epsilon}$  such that  $|\mathcal{A} + \mathcal{B}| \leq M^{1+\epsilon}$ , for some  $\mathcal{B}$  with  $|\mathcal{B}| \geq M^{1-10\epsilon}$ . Then  $\text{cp}(\mathcal{A} \cdot \mathcal{C})$  is smaller than  $\frac{1}{M^{1+10\epsilon}}$  for all sets  $\mathcal{C}$  of size  $M$ .*

We call a set  $\mathcal{A}$  satisfying the conclusion of Lemma 3.10 *times-friendly*. The main step in our proof will be the following lemma:

**Lemma 3.11.** *Let  $\mathcal{A} \subseteq \mathbb{F}$  with  $|\mathcal{A}| = M$ . Then, there exist two disjoint subsets  $\mathcal{A}_+$  and  $\mathcal{A}_\times$  such that*

- $\mathcal{A}_+$  is either empty or plus-friendly.
- $\mathcal{A}_\times$  is either empty or times-friendly.
- $|\mathcal{A} \setminus (\mathcal{A}_+ \cup \mathcal{A}_\times)| < M^{1-\epsilon}$ . (I.e.,  $\mathcal{A}_+ \cup \mathcal{A}_\times$  capture almost all of  $\mathcal{A}$ .)

Note that this lemma implies that the counterexample described above (of a distribution that is an arithmetic progression with probability half and a geometric progression with probability half) captures in some sense the worst case for the theorem.

*Proof of Lemma 3.11.* We prove the lemma by repeatedly applying Lemma 3.8 to construct the sets  $\mathcal{A}_+, \mathcal{A}_\times$ . We start with  $\mathcal{A}_+ = \emptyset$  and  $\mathcal{A}_\times = \mathcal{A}$ . At each point we remove some elements from  $\mathcal{A}_\times$  and add them to  $\mathcal{A}_+$ . We always maintain the invariant that  $\mathcal{A}_+$  is either empty or plus-friendly.

If  $|\mathcal{A}_\times| \leq M^{1-\epsilon}$  then we're done (since we can then let  $\mathcal{A}_\times = \emptyset$  and have  $|\mathcal{A}_\times \cup \mathcal{A}_+| \geq M - M^{1-\epsilon}$ ). If  $\mathcal{A}_\times$  is times-friendly then we're done. Otherwise, there exists  $\mathcal{B}''$  such that  $\text{cp}(\mathcal{A}_\times \cdot \mathcal{B}'') \geq \frac{1}{M^{1+\epsilon}}$  and so we can apply Lemma 3.8 to obtain subsets  $\mathcal{A}'$  of  $\mathcal{A}_\times$  and  $\mathcal{B}'$  of  $\mathcal{B}''$  such that  $|\mathcal{A}'|, |\mathcal{B}'| \geq M^{1-10\epsilon}$  but  $|\mathcal{A}' \cdot \mathcal{B}'| \leq M^{1+10\epsilon}$ . By Claim 3.9,  $\mathcal{A}'$  will be plus-friendly, so we can remove it from  $\mathcal{A}_\times$  and add it to  $\mathcal{A}_+$  (i.e., let  $\mathcal{A}_+ = \mathcal{A}_+ \cup \mathcal{A}'$ ,  $\mathcal{A}_\times = \mathcal{A}_\times \setminus \mathcal{A}'$ ). Note that the union of disjoint plus friendly sets is plus-friendly (since a convex combination of low collision-probability distributions has low collision probability by Lemma 3.5). We continue in this way until either  $\mathcal{A}_\times$  is times-friendly or  $|\mathcal{A}_\times| \leq M^{1-\epsilon}$ .  $\square$

<sup>19</sup>Let  $\mathcal{A}_{\text{arith}}$  denote the arithmetic progression. A random element from  $\mathcal{A} + \mathcal{A}$  is in the set  $\mathcal{A}_{\text{arith}} + \mathcal{A}_{\text{arith}}$  with probability at least  $\frac{1}{4}$ . However, because the set  $\mathcal{A}_{\text{arith}} + \mathcal{A}_{\text{arith}}$  is of size at most  $O(M)$ , we get that  $\text{cp}(\mathcal{A} + \mathcal{A}) \geq \frac{1}{16}\Omega(1/M) = \Omega(1/M)$ . The symmetrical reasoning shows that  $\text{cp}(\mathcal{A} \cdot \mathcal{A}) \geq \Omega(1/M)$ .

Using [Lemma 3.11](#), we can obtain a collision-probability analog of [Corollary 1.6](#):

**Lemma 3.12.** *Let  $\mathcal{A}, \mathcal{B}, \mathcal{C} \subseteq \mathbb{F}$  with  $|\mathcal{A}| = |\mathcal{B}| = |\mathcal{C}| = M$ . Then  $\mathcal{A} \cdot \mathcal{B} + \mathcal{C}$  is  $M^{-\epsilon}$ -close to having collision probability at most  $\frac{1}{M^{1+\epsilon}}$*

*Proof.* We split  $\mathcal{A}$  into  $\mathcal{A}_+$  and  $\mathcal{A}_\times$  as per [Lemma 3.11](#). The distribution  $\mathcal{A} \cdot \mathcal{B} + \mathcal{C}$  is within  $M^{-\epsilon}$  statistical distance to a convex combination of the distribution  $\mathcal{X}_+ = \mathcal{A}_+ \cdot \mathcal{B} + \mathcal{C}$  and the distribution of  $\mathcal{X}_\times = \mathcal{A}_\times \cdot \mathcal{B} + \mathcal{C}$ . (Unless  $\mathcal{A}_+$  or  $\mathcal{A}_\times$  is empty, in which case  $\mathcal{A} \cdot \mathcal{B} + \mathcal{C}$  is within  $M^{-\epsilon}$  statistical distance to one of these distributions.) We will finish the proof by showing that for both distributions  $\mathcal{X}_+$  and  $\mathcal{X}_\times$ , if the corresponding set is not empty, then the collision probability is at most  $\frac{1}{M^{1+\epsilon}}$ .

**Showing that  $\text{cp}(\mathcal{A}_+ \cdot \mathcal{B} + \mathcal{C}) \leq \frac{1}{M^{1+\epsilon}}$ .** Using [Lemma 3.5](#),  $\text{cp}(\mathcal{A}_+ \cdot \mathcal{B} + \mathcal{C}) \leq \max_{b \in \mathbb{F}} \text{cp}(\mathcal{A}_+ b + \mathcal{C})$ . However  $\text{cp}(\mathcal{A}_+ b + \mathcal{C}) = \text{cp}(\mathcal{A}_+ + \mathcal{C}b^{-1})$  since the latter distribution is a permutation of the former distribution (obtained by multiplying each element by  $b^{-1}$ ). Yet the fact that  $\mathcal{A}_+$  is plus-friendly implies that  $\text{cp}(\mathcal{A}_+ + \mathcal{C}b^{-1}) \leq \frac{1}{M^{1+\epsilon}}$ .

**Showing that  $\text{cp}(\mathcal{A}_\times \cdot \mathcal{B} + \mathcal{C}) \leq \frac{1}{M^{1+\epsilon}}$ .** This follows immediately from the fact that  $\text{cp}(\mathcal{A}_\times \cdot \mathcal{B}) \leq \frac{1}{M^{1+\epsilon}}$  and since  $\text{cp}(\mathcal{A}_\times \cdot \mathcal{B} + \mathcal{C}) \leq \text{cp}(\mathcal{A}_\times \cdot \mathcal{B})$  (as  $\mathcal{A}_\times \cdot \mathcal{B} + \mathcal{C}$  is a convex combination of distributions of the form  $\mathcal{A}_\times \cdot \mathcal{B} + c$  for some fixed  $c \in \mathcal{C}$ ).

□

### 3.3.1 Finishing up

[Lemma 3.12](#) almost directly implies [Lemma 3.1](#): First we use the known fact that if the distributions  $\mathcal{A}, \mathcal{B}, \mathcal{C}$  have min-entropy at least  $m$ , then the joint distribution  $\mathcal{A}, \mathcal{B}, \mathcal{C}$  is a convex combination on independent distributions of the form  $\mathcal{A}', \mathcal{B}', \mathcal{C}'$  where each of these is a uniform distribution on a set of size at least  $M = 2^m$ . Thus, it is sufficient to prove [Lemma 3.12](#) for such distributions. By [Lemma 3.12](#), for such distributions the distribution  $\mathcal{A}' \cdot \mathcal{B}' + \mathcal{C}'$  is within  $M^{-\epsilon}$  distance of having collision probability at most  $\frac{1}{M^{1+\epsilon}}$  for some constant  $\epsilon > 0$ . Now, by applying the [Lemma 3.6](#) (with  $K = M^{1+\epsilon/2}$ ,  $L = M^{\epsilon/2}$ ) we obtain that  $\mathcal{A}' \cdot \mathcal{B}' + \mathcal{C}'$  is within statistical distance  $M^{-\epsilon} + M^{-\epsilon/4}$  from having min-entropy at least  $\log(M^{(1+\epsilon/2)}) = (1 + \epsilon/2)m$ .

### 3.4 Constructing the field $\mathbb{F}$

Recall that our extractor obtains inputs in  $\{0, 1\}^n$ , and needs to treat these inputs as elements of a field  $\mathbb{F}$  of prime order. Unfortunately, there is no known deterministic polynomial-time algorithm that on input  $1^n$  outputs an  $n$ -bit prime (without assuming a strong number-theoretic assumption about the distance between consecutive primes). Fortunately, in our setting we can still find such a prime. The reason is that we can use some of our samples from the high-entropy distribution to obtain such a prime. To do so, we will use the following result on *seeded dispersers* (which we state here with the parameters suitable for our purposes):

**Theorem 3.13** ([\[Zuc91\]](#)). *For every  $\delta > 0$ , there exists a constant  $d > 1$  and a polynomial-time computable function  $D : \{0, 1\}^{(10/\delta)n} \times \{0, 1\}^{d \log n} \rightarrow \{0, 1\}^n$  such that for every set  $\mathcal{A} \subseteq \{0, 1\}^{(10/\delta)n}$  with  $|\mathcal{A}| \geq 2^{2n}$ , it holds that*

$$\left| D(\mathcal{A}, \{0, 1\}^{d \log n}) \cap \{0, 1\}^n \right| \geq \left(1 - \frac{1}{n^2}\right) \cdot 2^n$$

Let  $D$  be the function obtained from [Theorem 3.13](#) and identify its output with a number in  $[2^n]$ . We say that  $x \in \{0, 1\}^{10/\delta n}$  is “bad” if  $2^n + D(x, y)$  is *not* a prime number for *every*  $y \in \{0, 1\}^{d \log n}$ . Because the set of primes has at least  $1/n^2$  density in the interval  $[2^n, 2^{2n}]$ , the set of all bad  $x \in \{0, 1\}^{(10/\delta)n}$  is of size at most  $2^{2n}$ . This means that if we take  $10/\delta$  samples  $x_1, \dots, x_{10/\delta}$  from  $10/\delta$  independent distributions over  $\{0, 1\}^n$  each of min-entropy at least  $\delta n$ , then the probability that the concatenation  $x$  of  $x_1, \dots, x_{10/\delta}$  is bad is exponentially low (at most  $2^{2n}/2^{(10/\delta)\delta n} = 2^{-8n}$ ). This means that with  $1 - 2^{-\Omega(n)}$  probability if we enumerate over all seeds  $y \in \{0, 1\}^{d \log n}$  (which can be done in polynomial time), then we will find some  $y$  such that  $2^n + D(x, y)$  is prime. Note that we can check primality in deterministic polynomial-time [[AKS02](#)].<sup>20</sup> Thus we can construct the field  $\mathbb{F}$  with very high probability by using the first  $10/\delta n$  samples and the function  $D$  to obtain a prime  $P \in [2^n, 2^{n+1}]$ . We then embed the set  $\{0, 1\}^n$  in the field  $\mathbb{F} = \text{GF}(P)$ . Note that since  $n \geq \log P - 1$  there is no significant loss in relative entropy by this embedding.

**Using non-prime fields.** A different approach to solve the problem of obtaining the field  $\mathbb{F}$  is to use a *non-prime* field such as  $\text{GF}(2^p)$  (where  $p$  now is a *small* prime and hence can be easily found) and then use [Theorem 1.4](#) (the version proved in [Appendix A](#)) instead of [Theorem 1.5](#). This would yield the following variant of [Lemma 3.1](#):

**Lemma 3.14.** *There exists an absolute constant  $c > 0$  such that for every prime  $p$  every  $\delta > 0$  and every distributions  $\mathcal{A}, \mathcal{B}, \mathcal{C}$  over  $\text{GF}(2^p)$  with  $H^\infty(\mathcal{A}), H^\infty(\mathcal{B}), H^\infty(\mathcal{C}) > \delta p$ , the distribution  $\mathcal{A} \cdot \mathcal{B} + \mathcal{C}$  is  $2^{-\epsilon m}$ -close to having min-entropy at least  $\min\{(1 + \epsilon)m, 0.9 \log |\mathbb{F}|\}$  where  $\epsilon = c\delta$ .*

### 3.5 Decreasing the statistical distance.

We now show how we can decrease the statistical distance by repetition. We use the following variant of a XOR-lemma:

**Lemma 3.15.** *Let  $\mathcal{Y}_1, \dots, \mathcal{Y}_t$  be independent distributions over  $\mathbb{F}$  such that  $\text{dist}(\mathcal{Y}_i, U_{\mathbb{F}}) < \epsilon$  for every  $i = 1, \dots, t$ . Then*

$$\text{dist}(\mathcal{Y}_1 + \dots + \mathcal{Y}_t, U_{\mathbb{F}}) \leq \epsilon^t$$

*Proof.* We can represent each distribution  $\mathcal{Y}_i$  as a convex combination of the following form: with probability  $(1 - \epsilon)$  we get  $U_i$  (where each  $U_i$  is an independent copy of the uniform distribution) and with probability  $\epsilon$  an element of some distribution  $\tilde{\mathcal{Y}}_i$ . Thus, one can think of the distribution  $\mathcal{Y}_1 + \dots + \mathcal{Y}_t$  as a convex combination where with probability  $\epsilon^t$  we get an element of the form  $\tilde{\mathcal{Y}}_1 + \dots + \tilde{\mathcal{Y}}_t$  and with probability  $1 - \epsilon^t$  we get a distribution of the form  $Y + U_{\mathbb{F}}$  for some distribution  $Y$  which is equal to the uniform distribution.  $\square$

### 3.6 Proof of [Lemma 3.2](#)

Before proving [Lemma 3.2](#), we’ll prove the following related lemma:

**Lemma 3.16.** *Let  $F$  be any field of size  $N$ , and let  $\mathcal{A}, \mathcal{B}, \mathcal{C}, \mathcal{D}$  be four independent random variables over  $\mathbb{F}$ , where each variable has collision probability at most  $1/M$  for some  $M > |N|^{3/4}$ . Then,*

$$\text{dist}((\mathcal{A} - \mathcal{B}) \cdot (\mathcal{C} - \mathcal{D}), U_{\mathbb{F}}) < O(N^{3/2}/M^2)$$

where  $U_{\mathbb{F}}$  denotes the uniform distribution over  $\mathbb{F}$ .

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<sup>20</sup>It is possible to use the same idea to run also a *probabilistic* primality testing algorithm using some additional samples from the high entropy sources.

We will start with a variant, where we divide rather than multiply, and where  $A = C$  and  $B = D$ .

**Lemma 3.17.** *Let  $\mathcal{A}_1, \mathcal{A}_2$  be identical independent random variables over  $F$  with collision probability at most  $1/M$  for some  $M > N^{1/2}$  and let  $\mathcal{B}_1, \mathcal{B}_2$  be likewise. Then,*

$$\text{dist}((\mathcal{A}_1 - \mathcal{A}_2) \cdot (\mathcal{B}_1 - \mathcal{B}_2)^{-1}, U_{\mathbb{F}}) \leq O(N/M^2)$$

*Proof.* Let  $s$  be an arbitrary non-zero element of  $\mathbb{F}$ . The distribution  $\mathcal{A}_1 + s\mathcal{B}_1$  is a random variable over  $\mathbb{F}$  and hence has collision probability at least  $\frac{1}{N}$ . Hence we see that

$$\Pr[(\mathcal{A}_1 + s\mathcal{B}_1) = \mathcal{A}_2 + s\mathcal{B}_2] \geq \frac{1}{N}$$

. There are two ways that this equality can occur. If  $\mathcal{B}_1 = \mathcal{B}_2$  then this equation can only hold if  $\mathcal{A}_1 = \mathcal{A}_2$ . Otherwise, it holds only if  $s = (\mathcal{A}_1 - \mathcal{A}_2) \cdot (\mathcal{B}_1 - \mathcal{B}_2)^{-1}$  and hence we get that

$$\Pr[s = (\mathcal{A}_1 - \mathcal{A}_2) \cdot (\mathcal{B}_1 - \mathcal{B}_2)^{-1}] + \Pr[\mathcal{A}_{(1)} = \mathcal{A}_{(2)} \wedge \mathcal{B}_{(1)} = \mathcal{B}_{(2)}] \geq \frac{1}{N}$$

So  $\Pr[s = (\mathcal{A}_1 - \mathcal{A}_2) \cdot (\mathcal{B}_1 - \mathcal{B}_2)^{-1}] \geq \frac{1}{N} - 1/M^2$  for all  $s \neq 0$ , so almost all field elements are at least close to the uniform probability. The distance between two distributions  $D_1$  and  $D_2$  can be expressed as  $2 \sum_{s | D_1(s) \geq D_2(s)} (D_1(s) - D_2(s))$ . Taking  $D_1$  as the uniform distribution on  $F$ , and  $D_2$  as the above distribution, each  $s$  except 0 contributes at most  $1/M^2$ , and 0 contributes at most  $1/N$  to this sum. Hence, the statistical distance is at most  $2N/M^2 + 2/N = O(N/M^2)$ .  $\square$

**Lemma 3.18.** *Let  $X$  and  $Y$  be independent random variables on a set  $F$  of size  $N$ . Then  $\text{cp}(X) + \text{cp}(Y) + 2\Pr[X = Y] \geq 4/N$ .*

*Proof.* Let  $X$  be distributed according to  $D_0$  and  $Y$  according to  $D_1$ . Then let  $D_{1/2} = 1/2D_0 + 1/2D_1$ . Then  $D_{1/2}$  is a probability distribution on  $F$ , so  $\text{cp}(D_{1/2}) \geq 1/N$ . On the other hand,  $\text{cp}(D_{1/2}) = \sum_{s \in F} (1/2D_0(s) + 1/2D_1(s))^2 = 1/4(\sum_{s \in F} D_0(s)^2 + 2\sum_{s \in F} D_0(s)D_1(s) + \sum_{s \in F} D_1(s)^2) = 1/4(\text{cp}(X) + 2\Pr[X = Y] + \text{cp}(Y))$ .  $\square$

**Lemma 3.19.** *Let  $\mathcal{A}, \mathcal{B}, \mathcal{C}, \mathcal{D}$  be four independent variables over  $\mathbb{F}$  each with collision probability at most  $1/M$  where  $M > N^{1/2}$ . Let  $\mathcal{A}_1, \mathcal{A}_2$  be two independent variables distributed the same as  $\mathcal{A}$ , and likewise for  $\mathcal{B}_1, \mathcal{B}_2, \mathcal{C}_1, \mathcal{C}_2, \mathcal{D}_1, \mathcal{D}_2$ . Then for any  $s \in F$ ,  $s \neq 0$ ,  $\Pr[s = (\mathcal{A}_1 - \mathcal{A}_2)(\mathcal{B}_2 - \mathcal{B}_1)^{-1}] + \Pr[s = (\mathcal{C}_1 - \mathcal{C}_2)(\mathcal{D}_2 - \mathcal{D}_1)^{-1}] + 2\Pr[s = (\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B})^{-1}] \geq 4/N - O(1/M^2)$ .*

*Proof.* Let  $X$  be  $\mathcal{A} + s\mathcal{B}$  and  $Y$  be  $\mathcal{C} + s\mathcal{D}$ . Note that, by a similar case analysis to the proof of [Lemma 3.17](#),  $\text{cp}(X) \leq \Pr[s = (\mathcal{A}_1 - \mathcal{A}_2)(\mathcal{B}_2 - \mathcal{B}_1)^{-1}] + 1/M^2$ ,  $\text{cp}(Y) \leq \Pr[s = (\mathcal{C}_1 - \mathcal{C}_2)(\mathcal{D}_2 - \mathcal{D}_1)^{-1}] + 1/M^2$  and  $\Pr[X = Y] \leq \Pr[s = (\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B})^{-1}] + 1/M^2$  (using [Lemma 3.3](#) to bound the probabilities that  $\mathcal{A} = \mathcal{C}$  and  $\mathcal{B} = \mathcal{D}$ ). The claim then follows from [Lemma 3.18](#).  $\square$

**Lemma 3.20.** *Let  $\mathcal{A}, \mathcal{B}, \mathcal{C}, \mathcal{D}$  be four independent random variables over  $\mathbb{F}$  each with collision probability at most  $1/M$  where  $M > N^{1/2}$ . Then,*

$$\text{dist}((\mathcal{A} - \mathcal{B}) \cdot (\mathcal{C} - \mathcal{D})^{-1}, U_{\mathbb{F}}) < O(N/M^2)$$

where  $U_{\mathbb{F}}$  denotes the uniform distribution over  $\mathbb{F}$ .

*Proof.* Using [Lemma 3.19](#), we see that for all  $s \neq 0$ , with  $\Pr[s = (\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B})^{-1}] \leq 1/N$  the difference between this probability and  $1/N$  is at most  $O(1/M^2) + 1/2|1/N - \Pr[s = (\mathcal{A}_1 - \mathcal{A}_2)(\mathcal{B}_2 - \mathcal{B}_1)^{-1}]| + 1/2|1/N - \Pr[s = (\mathcal{C}_1 - \mathcal{C}_2)(\mathcal{D}_2 - \mathcal{D}_1)^{-1}]|$ . Thus, the statistical difference in the case of four distinct distributions is at most  $O(N/M^2)$  more than that when there are only two distinct distributions, which by [Lemma 3.17](#) is also  $O(N/M^2)$ .  $\square$

**Lemma 3.21.** *Let  $\mathcal{Z}$  be a random variable over  $F$ . Let  $\delta = \text{dist}(\mathcal{Z}, U_F)$ . Then  $1/N + \delta^2/N \leq \text{cp}(\mathcal{Z}) \leq 1/N + \delta^2$ .*

*Proof.* Let  $\delta_s = |\Pr[\mathcal{Z} = s] - 1/N|$ . Then  $\sum_{s \in F} \delta_s = \delta$ , and  $\sum_{s \in F} \delta_s^2 = \sum_s (\Pr[\mathcal{Z} = s])^2 - 2/N \sum_s [\Pr[\mathcal{Z} = s] + N/N^2] = \text{cp}(\mathcal{Z}) - 1/N$ . By convexity,  $\sum \delta_s^2$  is minimized when all  $\delta_s = \delta/N$  and maximized when one  $\delta_s = \delta$ . Thus,  $\delta^2/N \leq \text{cp}(\mathcal{Z}) - 1/N \leq \delta^2$ .  $\square$

To obtain [Lemma 3.16](#), we note that for every two distributions  $\mathcal{X}, \mathcal{Y}$  over  $\mathbb{F} \setminus \{0\}$ ,  $\text{cp}(\mathcal{X} \cdot \mathcal{Y}) = \text{cp}(\frac{\mathcal{X}}{\mathcal{Y}})$ . Indeed,  $x \cdot y = x' \cdot y'$  if and only if  $\frac{x}{y} = \frac{x'}{y'}$ . By [Lemma 3.20](#),  $\text{dist}((\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B})^{-1}, \mathcal{U}) = O(N/M^2)$ . Thus, from the previous lemma,  $\text{cp}((\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B})) = \text{cp}((\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B})^{-1}) \leq 1/N + O(N^2/M^4)$ . Then it follows from the lemma above that  $\text{dist}((\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B}), \mathcal{U})^2/N \leq O(N^2/M^4)$ , so  $\text{dist}((\mathcal{A} - \mathcal{C})(\mathcal{D} - \mathcal{B}), \mathcal{U}) \leq O(N^{3/2}/M^2)$ .  $\square$

**Proving Lemma 3.2.** We can now prove [Lemma 3.2](#). Indeed, for every 9 distributions  $\mathcal{X}_1, \dots, \mathcal{X}_9$ ,  $\text{cp}(\text{Ext}^2(\mathcal{X}_1, \dots, \mathcal{X}_9)) \leq \text{cp}(\text{Ext}^1(\mathcal{X}_1, \mathcal{X}_2, \mathcal{X}_3) \cdot \text{Ext}^1(\mathcal{X}_4, \mathcal{X}_5, \mathcal{X}_6))$ , since adding the additional independent distribution  $\text{Ext}^1(\mathcal{X}_7, \mathcal{X}_8, \mathcal{X}_9)$  cannot increase the collision probability. Hence, it is enough to prove that for  $\mathcal{X}_1, \dots, \mathcal{X}_6$  of min-entropy  $0.9 \log N$ , the distribution  $(\mathcal{X}_1 \cdot \mathcal{X}_2 + \mathcal{X}_3) \cdot (\mathcal{X}_4 \cdot \mathcal{X}_5 + \mathcal{X}_6)$  is  $N^{-0.1}$  close to the uniform distribution over  $\mathbb{F}$ . Yet this distribution is a convex combination of distributions of the form  $(\mathcal{X}_1 \cdot x_2 + \mathcal{X}_3) \cdot (\mathcal{X}_4 \cdot x_5 + \mathcal{X}_6)$  for fixed (and with very high probability non-zero)  $x_2, x_5 \in \mathbb{F}$ . Any such distribution is of the form  $(\mathcal{A} - \mathcal{B})(\mathcal{C} - \mathcal{D})$  (for  $\mathcal{A} = \mathcal{X}_1 \cdot x_2$ ,  $\mathcal{B} = -\mathcal{X}_3$ ,  $\mathcal{C} = \mathcal{X}_4 \cdot x_5$ ,  $\mathcal{D} = -\mathcal{X}_6$ ) and hence the result is implied by [Lemma 3.16](#).  $\square$

## 4 A Constant-Samples Same-Source Disperser for Low Min-Entropy

In this section we prove [Theorem 1.3](#). That is, we construct a constant-samples same-source disperser for subsets of  $\{0, 1\}^n$  of size  $n^{O(1)}$ . A central tool we will use is the *deterministic coin tossing* technique. This technique was used in several contexts in computer science (e.g., the parallel algorithm of [\[CV86\]](#)), and it was also used in a very similar context to ours by Erdős and Hajnal (where it was called the “stepping up lemma”, see Section 4.7 in [\[GRS80\]](#)) and Fiat and Naor [\[FN93\]](#). By “deterministic coin tossing” we mean the function  $\text{ct} : \{0, 1\}^{2n} \rightarrow [n]$  defined as follows: for every  $x, y \in \{0, 1\}^n$ ,  $\text{ct}(x, y)$  is equal to the first position  $i$  such that  $x_i \neq y_i$  (if  $x = y$  then we let  $\text{ct}(x, y) = n$ ). The following property of this function will be useful for us:

**Lemma 4.1.** *For every  $A \subseteq \{0, 1\}^n$ ,  $|\text{ct}(A, A)| \geq \log |A|$*

*Proof.* Let  $S \subseteq [n]$  denote  $\text{ct}(A, A)$  and suppose for the sake of contradiction that  $|A| > 2^{|S|}$ . Then, by the pigeon-hole principle, there exist two distinct strings  $x, y \in A$  such that  $x$  agrees with  $y$  on all the coordinates of  $S$ . However, for such  $x$  and  $y$  clearly  $\text{ct}(x, y) \notin S$ .  $\square$

The main idea behind our construction of a disperser is to apply the function  $\text{ct}$  in order to map our initial domain  $\{0, 1\}^n$  to the much smaller domain  $[n]$ , and then use brute force to find an optimal disperser (or even extractor) on this smaller domain. In fact, we will need to apply the function *twice* to reduce our domain to the domain  $[\log n]$  so we can apply brute force and obtain an optimal disperser for the smaller domain. That is, we use the following simple lemma:

**Lemma 4.2.** *For every  $n, m$  (with  $m < n$ ) it is possible to construct in  $2^{O(n)}$  time the table for a function  $E_{\text{opt}} : [n] \times [n] \rightarrow [m]$  such that for every  $A \subseteq [n]$  with  $|A| > \log n + m + O(1)$ , it holds that  $|E_{\text{opt}}(A, A)| \geq \frac{m}{2}$*

*Proof Sketch:* One way to do so will be to go over all possible functions until we find a function satisfying this condition (there will be such a function because a random function satisfies it with high probability). However, enumerating all possible functions will take  $m^{n^2}$ -time. Note that testing this condition only requires going over all such subsets  $A$  which are at most  $\min\{2^n, n^{2m \log n}\}$  many. Thus if we reduce the number of functions to test then we can reduce the overall time to a polynomial in this number. This can be done by considering functions from a sample space which is  $\frac{1}{2m}$ -close to being 2 log  $n$ -wise independent. Such a sample space can be constructed in time  $2^{O(n)}$ .  $\square$

**Proof of Theorem 1.3.** To prove Theorem 1.3, we consider the following disperser  $\text{Disp}$ :

$$\text{Disp}(x_1, \dots, x_8) = E_{\text{opt}}(\text{ct}(\text{ct}(x_1, x_2), \text{ct}(x_3, x_4)), \text{ct}(\text{ct}(x_5, x_6), \text{ct}(x_7, x_8)))$$

It clearly runs in polynomial time. We will prove that for every set  $A$  of size at least  $n^{5m}$ ,  $|\text{Disp}(A, \dots, A)| \geq \frac{m}{2}$ . Thus  $\text{Disp}$  is a disperser with “statistical distance” equal to  $\frac{1}{2}$ . Such a disperser can be modified to obtain a disperser according to our standard definition. For example, this can be done by embedding  $[m]$  in some prime field  $[p]$  (with  $m \leq p \leq 2m$ ) and letting  $\text{Disp}'(x_1, \dots, x_{32}) = \text{Disp}(x_1, \dots, x_8) + \text{Disp}(x_9, \dots, x_{16}) + \text{Disp}(x_{17}, \dots, x_{24}) + \text{Disp}(x_{25}, \dots, x_{32})$ . By the Cauchy-Davenport inequalities it will hold that for every set  $A$  as above,  $\text{Disp}'(A, \dots, A) = [p]$ .

To prove the bound on  $A$ , we note that by Lemma 4.1, if  $|A| \geq n^{5m}$  then  $\text{ct}(\text{ct}(A, A), \text{ct}(A, A)) \geq \log \log n + m + 5$ . This means that  $|E_{\text{opt}}(A, A)| \geq \frac{m}{2}$ , thus finishing our proof.  $\square$

## 5 Subsequent and Future Work

In this work we have given the first extractors for a constant number of independent sources, each of linear entropy. Unfortunately the number of sources needed is a function of the (constant) entropy rate, and the most obvious next challenge was to remove this dependency.

In very recent work, [BKS<sup>+</sup>05] have achieved this. They give explicit deterministic extractors from 3 independent sources of any linear entropy. This was further improved by Ran Raz [Raz05], who only needs one of the sources to have linear entropy, while the others can have only logarithmic amount of entropy.

Further results in [BKS<sup>+</sup>05] give a deterministic disperser from only two independent sources of constant entropy rate (greatly improving bounds on the Bipartite Ramsey problem). They also give a disperser for one *affine* source (a uniform distribution of an affine subspace) of linear entropy (=dimension). The error in all these constructions is a constant.

All these new results heavily rely on the techniques and results in this paper. They demonstrate the power of the tools from combinatorial number theory even more, and argue the relevance of these tools to theoretical computer science.

In light of this progress, the two natural directions which present themselves. The first is reducing the entropy requirement in extraction from a constant number of sources. Can we go down to polynomial, or even polylogarithmic entropy? The second is reducing the error in all these constructions. Can we go down to polynomial, or even exponential error? The first problem seems to require a way around the [BKT03] argument (which necessarily needs linear entropy). The second problem seem to require bypassing the use of a non-constructive extractor as a building block. Both need new ideas!

Another major problem is of course extraction from two sources. The dispersers in [BKS<sup>+</sup>05] are strong in the sense that each of their outputs has some constant probability. But the technique

achieving that, called the challenge-response mechanism in that paper, seem to inherently prevent extraction. Can one bypass this method, or enhance it? The same question applies equally well to the problem of extraction from affine sources. An exciting new development in this direction is a very recent result by Jean Bourgain [Bou05], which gives (again using results from [BKT03]) a simple algebraic construction of an extractor for 2 independent sources requiring  $(\frac{1}{2} - \epsilon)n$  entropy for some small absolute constant  $\epsilon > 0$ .

To summarize, after nearly 20 years of no progress, there seem to be new excitement, ideas, tools, and results on deterministic extraction from independent sources, and we expect more to follow soon.

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## A A Proof of [Theorem 1.4](#)

In this section, we prove [Theorem 1.4](#) with a better explicit dependence between the constants  $\epsilon$  and  $\delta$  (namely  $\epsilon = \Theta(\delta)$ ). That is, we prove the following theorem:

**Theorem A.1.** *There exists an absolute constant  $c_0 > 0$  such that for every field  $\mathbb{F}$  and  $\delta > 0$  such that  $|\mathbb{F}|^{1/k}$  is not an integer for every  $k \in [2 \dots (2/\delta)]$ , and every set  $\mathcal{A} \subseteq \mathbb{F}$  with  $|\mathbb{F}|^\delta < |\mathcal{A}| < |\mathbb{F}|^{1-\delta}$ , it holds that*

$$\max\{|\mathcal{A} + \mathcal{A}|, |\mathcal{A} \cdot \mathcal{A}|\} \geq |\mathcal{A}|^{1+c_0\delta}$$

Note that [Theorem A.1](#) indeed implies [Theorem 1.4](#) since a finite field  $\mathbb{F}$  has a subfield of size  $M$  if and only if  $M = |\mathbb{F}|^{1/k}$  for some integer  $k$ .

We prove [Theorem A.1](#) by proving the following two claims:

**Claim A.2.** *There exists a fixed size rational expression  $r(\cdot)$  such that for every  $\delta$ , field  $\mathbb{F}$  satisfying the conditions of [Theorem A.1](#) and set  $\mathcal{B} \subseteq \mathbb{F}$  with  $|\mathcal{B}| \geq |\mathbb{F}|^\delta$ ,*

$$|r(\mathcal{B}, \dots, \mathcal{B})| \geq \min\{|\mathcal{B}|^{1+\delta}, |\mathbb{F}|\}$$

By a *rational expression* we mean an expression involving only the operations  $+$ ,  $-$ ,  $\cdot$  and  $/$  and variable names, but no coefficients (for example,  $(x_1 - x_2)/(x_2 \cdot x_3 + x_4)$ ). By *fixed size* we mean that the number of operations and variables in the expression does not depend on  $\delta$ ,  $\mathbb{F}$  or  $\mathcal{A}$ . In fact, the expression  $r$  obtained from the proof of [Claim A.2](#) will be very simple: it is a 16-variable rational expression with both the nominator and the denominator -of degree 2.

**Claim A.3.** *For every rational expression  $r(\cdot)$  there is a constant  $c$  (depending on  $r(\cdot)$ ) such that if  $\mathcal{A} \subseteq \mathbb{F}$  satisfies  $|\mathcal{A} + \mathcal{A}|, |\mathcal{A} \cdot \mathcal{A}| \leq |\mathcal{A}|^{1+\rho}$  (for sufficiently small  $\rho > 0$ ) then there exists a set  $\mathcal{B} \subseteq \mathcal{A}$  with  $|\mathcal{B}| \geq |\mathcal{A}|^{1-c\rho}$  but  $|r(\mathcal{B}, \dots, \mathcal{B})| \leq |\mathcal{A}|^{1+c\rho}$ .*

**Proving [Theorem A.1](#) using [Claims A.2](#) and [A.3](#).** [Claims A.2](#) and [A.3](#) together imply [Theorem A.1](#). Indeed let  $r(\cdot)$  be the rational expression obtained from [Claim A.2](#) and let  $c$  be the constant (depending on  $r$ ) obtained from [Claim A.3](#). If for  $\delta, \mathbb{F}$  and  $\mathcal{A}$  as in the conditions of the theorem both  $|\mathcal{A} + \mathcal{A}|$  and  $|\mathcal{A} \cdot \mathcal{A}|$  are less than  $|\mathcal{A}|^{1+\delta/(10c)}$  then there is a subset  $\mathcal{B}$  of size more than  $|\mathcal{A}|^{1-c\delta/(10c)} \geq |\mathbb{F}|^{\delta/2}$  such that  $|r(\mathcal{B}, \dots, \mathcal{B})| \leq |\mathcal{A}|^{1+\delta/10} < |\mathcal{B}|^{1+\delta/2}$ , thus obtaining a contradiction.  $\square$

We note that [Claim A.3](#) follows almost immediately from Lemmas 2.4 and 3.1 in [\[BKT03\]](#). Nevertheless, for the sake of completeness, we do provide a proof of [Claim A.3](#) (following similar lines to the proofs of [\[BKT03\]](#)) in [Appendix A.2](#). We now move to the proof of [Claim A.2](#).

## A.1 Proof of [Claim A.2](#)

To prove [Claim A.2](#) we'll prove the following even simpler claim:

**Claim A.4.** *Let  $\mathbb{F}$  be any field and let  $\mathcal{B} \subseteq \mathbb{F}$  and  $k \in \mathbb{N}$  (with  $k \geq 2$ ) be such that  $|\mathbb{F}|^{1/k} < |\mathcal{B}| \leq |\mathbb{F}|^{1/(k-1)}$ . Then  $|\frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}| \geq |\mathbb{F}|^{1/(k-1)}$ .*

*Proof.* Suppose otherwise that  $|\frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}| < |\mathbb{F}|^{1/(k-1)}$ . Thus we can find  $s_1 \notin \frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}$ . Similarly, if  $k > 2$  then we can find  $s_2 \notin \frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}} + s_1 \frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}$  (since this set is of size at most  $|\frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}|^2 < |\mathbb{F}|$ ). In this way we define inductively  $s_1, \dots, s_{k-1}$  such that

$$s_i \notin \frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}} + s_1 \frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}} + \dots + s_{i-1} \frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}$$

Consider the function  $f(x_0, \dots, x_{k-1}) = x_0 + s_1 x_1 + \dots + s_{k-1} x_{k-1}$ . This is a function from  $\mathcal{B}^k$  to  $\mathbb{F}$  where  $|\mathcal{B}|^k > |\mathbb{F}|$  and hence it has a *collision*. That is, there are two vectors  $\vec{x} = (x_0, \dots, x_{k-1})$  and  $\vec{x}' = (x'_0, \dots, x'_{k-1})$  such that  $\vec{x} \neq \vec{x}'$  but  $f(\vec{x}) = f(\vec{x}')$ . If we let  $i$  be the maximum index such that  $x_i \neq x'_i$  we see that

$$(x_0 - x'_0) + s_1(x_1 - x'_1) + \dots + s_{i-1}(x_{i-1} - x'_{i-1}) = s_i(x'_i - x_i) \quad .$$

Dividing by  $(x'_i - x_i)$  we get that  $s_i = y_0 + s_1 y_1 + \dots + s_{i-1} y_{i-1}$  where all the  $y_i$ 's are members of  $\frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}$ , contradicting our choice of  $s_i$ .  $\square$

To prove [Claim A.2](#) we let  $\mathbb{F}, \delta, \mathcal{B}$  be as stated in the theorem and choose  $k$  such that  $|\mathbb{F}|^{1/k} < |\mathcal{B}| \leq |\mathbb{F}|^{1/(k-1)}$ . By one invocation of  $\frac{\mathcal{B}-\mathcal{B}}{\mathcal{B}-\mathcal{B}}$  we get to a set of size at least  $|\mathbb{F}|^{1/(k-1)}$  but since that is not an integer, this set is of size *larger* than  $|\mathbb{F}|^{1/(k-1)}$ . Thus if we compose this expression two times (i.e., let  $r(x_1, \dots, x_{16}) = \frac{r'(x_1, \dots, x_4) - r'(x_5, \dots, x_8)}{r'(x_9, \dots, x_{12}) - r'(x_{13}, \dots, x_{16})}$  where  $r'(a, b, c, d) = \frac{a-b}{c-d}$ ) then we get that

$$|r(\mathcal{B}, \dots, \mathcal{B})| \geq |\mathbb{F}|^{\frac{1}{k-2}} = |\mathbb{F}|^{\frac{1}{k-1}(1 + \frac{1}{k-2})} \geq |\mathcal{B}|^{1+\delta}$$

□

## A.2 Proof of [Claim A.3](#)

We now prove [Claim A.3](#). Before turning to the actual proof, we state two number theoretic lemmas which we'll use. These lemmas are variants of the lemmas presented in [Section 3.2](#).

### A.2.1 More number theoretic lemmas.

**Stronger form of Gowers's Lemma.** We will use the following generalized and stronger form of [Lemma 3.8](#) (see [[Gow98](#), [BKT03](#)], [[SSV04](#), Claim 4.4]). Because it is such a useful lemma, we state it below in the most general (and unfortunately also cumbersome) form:

**Lemma A.5.** *Let  $\mathcal{A}_1, \dots, \mathcal{A}_k$  be subsets of some group  $G$  with  $|\mathcal{A}_i| = M$  for all  $i \in [k]$ . Then, there exists  $C = C(k)$  such that for every  $\rho > 0$ , if  $\text{cp}(\sum_{i=1}^k \mathcal{A}_i) \geq \frac{1}{M^{1+\rho}}$  then there are  $k$  subsets  $\mathcal{A}'_1, \dots, \mathcal{A}'_k$  with  $\mathcal{A}'_i \subseteq \mathcal{A}_i$  and  $|\mathcal{A}'_i| \geq M^{1-C\rho}$  for every  $i \in [k]$  satisfying*

$$|\sum_{i=1}^k \mathcal{A}'_i| \leq M^{1+C\rho} \tag{1}$$

Moreover, [Equation 1](#) is demonstrated by the fact that every element  $z \in \sum_{i=1}^k \mathcal{A}'_i$  can be represented in  $M^{\ell-1-C'\rho}$  different ways as a sum  $z = y_1 + \dots + y_\ell$  (for  $C' = C'(k)$  and  $\ell = 2k^2 - k$ ) where each of the  $y_j$ 's is a member of  $\mathcal{A}_i$  or  $-\mathcal{A}_i$  for some  $i = i(j)$  (with the choice of a sign also being a function of  $j$ ).

Furthermore, if for all  $i \in [k]$ ,  $\mathcal{A}_i = \mathcal{A}$  or  $\mathcal{A}_i = -\mathcal{A}$  for some set  $\mathcal{A}$  then all the subsets  $\mathcal{A}'_i$  are of the form  $\mathcal{A}'_i = \mathcal{A}'$  or  $\mathcal{A}'_i = -\mathcal{A}'$  for some subset  $\mathcal{A}' \subseteq \mathcal{A}$ .

It is easy to see that by applying [Lemma A.5](#) for  $k = 2$  we get [Lemma 3.8](#) (perhaps with 10 replaced by a different constant). [Lemma A.5](#) can be proven by a generalization of the proof of [Lemma 3.8](#), see [[SSV04](#)].<sup>21</sup> When using [Lemma A.5](#), we will always be in the case that that we have an upper bound on the *set size* of  $\sum_{i=1}^k \mathcal{A}_i$  (i.e.,  $|\sum_{i=1}^k \mathcal{A}_i| \leq M^{1+\rho}$ ) and not just a lower bound on its collision probability.

We note that in the proofs below we will use several times this technique of showing that some set  $\mathcal{B}$  is not much larger than  $M$  by showing that every  $b \in \mathcal{B}$  can be represented in roughly  $M^{\ell-1}$  ways as a sum of  $\ell$  elements, each from some set  $\mathcal{D}$  of size  $M$ .

<sup>21</sup>We note that this lemma is not stated exactly in this form in [[SSV04](#)]. Rather, Claim 4.4 there states that under these conditions every such  $z$  can be represented in roughly  $M^{2k-2}$  ways as a sum of  $2k-1$  elements  $w_i$ , where each of these elements can be represented in roughly  $M^{k-1}$  ways as a sum of  $k$  elements in the  $\mathcal{A}_i$ 's.

**Sunset estimates.** We'll also use the following Lemma, which is a variant of [Lemma 3.7](#):

**Lemma A.6** ([\[Plü70, Ruz96\]](#)). *Let  $\mathcal{A}, \mathcal{B}$  be subsets of some Abelian group  $G$  with  $|\mathcal{A}| = |\mathcal{B}| = M$  and let  $\rho > 0$  be some number. If  $|\mathcal{A} + \mathcal{B}| \leq M^{1+\rho}$  then*

$$|\mathcal{A} \pm \underbrace{\mathcal{B} \cdots \pm \mathcal{B}}_{h \text{ times}}| \leq M^{1+2h\rho}$$

### A.2.2 The actual proof.

In fact, to prove [Claim A.3](#) it is enough to prove the following claim:

**Claim A.7.** *For every integer  $k > 0$  there exists a constant  $C = C(k) > 0$  such that for every  $\rho > 0$ , if  $\mathcal{A} \subseteq \mathbb{F}$  satisfies  $|\mathcal{A} + \mathcal{A}|, |\mathcal{A} \cdot \mathcal{A}| \leq |\mathcal{A}|^{1+\rho}$  then there is a set  $\mathcal{B} \subseteq \mathcal{A}$  such that  $|\mathcal{B}| \geq |\mathcal{A}|^{1-C\rho}$  but  $|\mathcal{B}^k - \mathcal{B}^k| \leq |\mathcal{A}|^{1+C\rho}$ .*

**Obtaining [Claim A.3](#) from [Claim A.7](#).** [Claim A.7](#) implies [Claim A.3](#) by applying [Lemma A.6](#). Indeed, suppose that  $r$  is a rational expression where both the numerator and denominator have at most  $k'$  monomials each of degree at most  $k'$ . Let  $k = 4k'^2$  and let  $\mathcal{B}$  be the subset of  $\mathcal{A}$  obtained from [Claim A.7](#) such that  $|\mathcal{B}^k - \mathcal{B}^k|$  is at most  $|\mathcal{A}|^{1+C\rho}$  for some constant  $C = C(k)$ . Since  $|\mathcal{B}| \geq |\mathcal{A}|^{1-C\rho}$ , this means that  $|\mathcal{B}^k - \mathcal{B}^k| \leq |\mathcal{B}|^{1+C'\rho}$  for some different constant  $C'$ . This implies by [Lemma 3.7](#) that  $|k\mathcal{B}^k| \leq |\mathcal{B}|^{1+2C'k\rho}$ . Now,  $k\mathcal{B}^k \supseteq \mathcal{C} \stackrel{\text{def}}{=} (k'\mathcal{B}^{k'})(k'\mathcal{B}^{k'})$  and hence the size of  $\mathcal{C}$  is also at most  $|\mathcal{B}|^{1+2C'k\rho}$ . Applying [Lemma 3.7](#) again we get that  $\left| \frac{k'\mathcal{B}^{k'}}{k'\mathcal{B}^{k'}} \right| \leq |\mathcal{A}|^{1+4C'k\rho}$ . However,  $r(\mathcal{B}, \dots, \mathcal{B}) \subseteq \frac{k'\mathcal{B}^{k'}}{k'\mathcal{B}^{k'}}$  and hence we're done.

*Proof of [Claim A.7](#).* We now turn to proving [Claim A.7](#). Let  $\mathcal{A}$  be a set satisfying the conditions of the claim, and denote  $M = |\mathcal{A}|$ . We will prove that for some constant  $C = C(k) > 0$ , there is a subset  $\mathcal{B} \subseteq \mathcal{A}$  of size at least  $M^{1-C\rho}$  such that any member of the set  $\mathcal{B}^k - \mathcal{B}^k$  can be represented in at least  $M^{\ell-1-C\rho}$  different ways as a sum  $d_1 + \dots + d_\ell$  where all the elements  $d_i$  come from a set  $\mathcal{D}$  of size at most  $M^{1+C\rho}$  for some  $\ell = \ell(k)$ . This clearly implies that  $|\mathcal{B}^k - \mathcal{B}^k| \leq M^{1+C'\rho}$  for some constant  $C'$  thus proving the claim.

**Proof idea: the case  $k = 2$ .** To illustrate the proof idea, we now sketch the proof for the case  $k = 2$ . The proof for general  $k$  follows in exactly the same way, although with more cumbersome notations. By [Lemma A.5](#), under the conditions of the claim, there exists a set  $\mathcal{A}' \subseteq \mathcal{A}$  with  $|\mathcal{A}'| \geq M^{1-C\rho}$  (where  $C$  is some absolute constant independent of  $\rho$ ) such that not only  $|\mathcal{A}' - \mathcal{A}'| \leq M^{1+C\rho}$  but actually this can be demonstrated by the fact that every element of  $\mathcal{A}' - \mathcal{A}'$  can be represented in at least  $M^{5-C\rho}$  different ways in the form  $a_1 - a_2 + a_3 - a_4 + a_5 - a_6$  where for  $i = 1, \dots, 6$ ,  $a_i \in \mathcal{A}$ .<sup>22</sup>

We know that every member of  $\mathcal{A}' - \mathcal{A}'$  can be represented in roughly  $M^5$  different ways as  $a_1 - a_2 + a_3 - a_4 + a_5 - a_6$  with the  $a_i$ 's in  $\mathcal{A}$ . Now, if we multiply this by an arbitrary element of  $\mathcal{A}$  we get that every member of  $(\mathcal{A}' - \mathcal{A}')\mathcal{A}$  can be represented in roughly  $M^5$  different ways as  $b_1 - b_2 + b_3 - b_4 + b_5 - b_6$  with the  $b_i$ 's in  $\mathcal{A} \cdot \mathcal{A}$ . Since by the conditions of the claim,  $\mathcal{A} \cdot \mathcal{A}$  is also of roughly size  $M$ , we get that the set  $(\mathcal{A}' - \mathcal{A}')\mathcal{A}$  is also “not large” (i.e., of size  $M^{1+C'\rho}$  for some absolute constant  $C'$ ). Now consider an element  $y - z$  of the set  $\mathcal{A}'^2 - \mathcal{A}'^2$ . For the sake of

<sup>22</sup>For simplicity of exposition we assumed that the pattern of + and - signs that is obtained from [Lemma A.5](#) is as written above. The proof clearly follows through regardless of the fixed pattern we use.

simplicity, we assume for a moment (*with loss of generality*) that any  $y$  in  $\mathcal{A}' \cdot \mathcal{A}'$  can be represented in roughly  $M$  different ways as  $y = y_1 y_2$  with  $y_1, y_2 \in \mathcal{A}$ .<sup>23</sup> Since  $z \in \mathcal{A}' \cdot \mathcal{A}'$ , it is equal to  $z_1 z_2$  for some  $z_1, z_2 \in \mathcal{A}$ . Every representation  $y_1 y_2$  of  $y$  induces a representation of  $y - z = y_1 y_2 - z_1 z_2$  as  $(y_1 - z_1) z_2 + z_1 (y_2 - z_2)$ , and so we get that every element of  $\mathcal{A}'^2 - \mathcal{A}'^2$  can be represented as  $d_1 - d_2$ , with  $d_1, d_2 \in \mathcal{D} \stackrel{\text{def}}{=} (\mathcal{A}' - \mathcal{A}') \mathcal{A}$  in roughly  $M$  different ways. However, since we already showed that the size of  $\mathcal{D}$  is roughly  $M$ , this proves that  $\mathcal{A}'^2 - \mathcal{A}'^2$  is also of size roughly  $M$  and hence we're done (for the case  $k = 2$ ).

**Proving for general  $k$ .** We now turn to proving the claim for rigorously and for any  $k$ . Let  $\ell$  be some number, and let  $a, b \in \mathcal{A}'$  (where  $\mathcal{A}'$  is obtained from [Lemma A.5](#) as above) and  $c \in \mathcal{A}^\ell$ . By the same reasoning as above, we get that the element  $(a + b)c$  can be represented in at least  $M^{5-C\rho}$  different ways as  $a_1 c - a_2 c + a_3 c - a_4 c + a_5 c - a_6 c$  or in other words, it can be represented in at least  $M^{5-C\rho}$  different ways as  $b_1 - b_2 + b_3 - b_4 + b_5 - b_6$  for  $b_i \in \mathcal{A}^{\ell+1}$  for  $i = 1, \dots, 6$ . By the multiplicative version of [Lemma A.6](#),  $|\mathcal{A}^\ell| \leq M^{1+2\ell\rho}$  and hence the set  $(\mathcal{A}' - \mathcal{A}') \mathcal{A}^\ell$  is “small” (i.e., of size at most  $M^{1+4C(\ell)\rho}$ ). Using again the multiplicative version of [Lemma A.6](#) (setting  $\mathcal{A} = \mathcal{A}' - \mathcal{A}'$ ,  $\mathcal{B} = \mathcal{A}^\ell$ ), we get that the set  $(\mathcal{A}' - \mathcal{A}') \mathcal{A}^\ell \mathcal{A}'^{-\ell}$  is also “small” (i.e., of size at most  $M^{1+4C(\ell, \ell')\rho}$ ). Using the fact that  $(\mathcal{A}'^{-1} - \mathcal{A}'^{-1}) \subseteq (\mathcal{A}' - \mathcal{A}') \mathcal{A}'^{-2}$ , we get that for any  $\ell$ , the set  $\mathcal{D}_\ell = (\mathcal{A}'^{-1} - \mathcal{A}'^{-1}) \mathcal{A}^\ell \cup (\mathcal{A}' - \mathcal{A}') \mathcal{A}^\ell$  is of size at most  $M^{1+C'(\ell)\rho}$ . We define  $\mathcal{D} \stackrel{\text{def}}{=} \mathcal{D}_\ell$  for  $\ell$  as obtained by [Lemma A.5](#) (i.e.,  $\ell = 2k^2 - k$ ).

**Defining the set  $\mathcal{B}$ .** Utilizing [Lemma A.5](#) again, we obtain that for some absolute constant  $D = D(k)$ , there is a set  $\mathcal{B} \subseteq \mathcal{A}'$  with  $|\mathcal{B}| \geq M^{1-D\rho}$  and every element in  $\mathcal{B}^k$  can be represented in  $M^{\ell-1-D\rho}$  ways as  $a_1 \cdots a_\ell$ , where  $a_i$  is in  $\mathcal{A}$  or in  $\mathcal{A}^{-1}$  for  $i = 1, \dots, \ell$  (for  $\ell = 2k^2 - k$ ). Let  $y - z$  be a member of  $\mathcal{A}'^k - \mathcal{A}'^k$ . Fix one representation  $z = z_1 \cdots z_\ell$  of  $z$  as a multiplication of elements of  $\mathcal{A}$  or  $\mathcal{A}^{-1}$ . The element  $y$  can be represented in  $M^{\ell-1-D\rho}$  times as  $y = y_1 \cdots y_\ell$ . For each such representation we define  $d_i = y_1 \cdots y_{i-1} (y_i - z_i) z_{i+1} \cdots z_\ell$ . Note that  $d_i \in \mathcal{D}$ , where  $\mathcal{D}$  is the set defined above. The map  $(y_1, \dots, y_\ell) \mapsto (d_1, \dots, d_\ell)$  is one-to-one, and hence we get that each member of  $\mathcal{B}^k - \mathcal{B}^k$  can be represented in  $M^{\ell-1-D\rho}$  different ways as  $\sum_{i=1}^\ell d_i$  with  $d_i \in \mathcal{D}$ , implying that  $|\mathcal{B}^k - \mathcal{B}^k| < M^{1+D'\rho}$  for  $D' = D'(k)$ .  $\square$

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<sup>23</sup>We will not be able to get to that situation in the actual proof below, but we will approximate it using the multiplicative version of [Lemma A.5](#).